# The non-preemptive 'Join the Shortest Queue-Serve the Longest Queue' service system with or without switch-over times 

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#### Abstract

A 2-queue system with a single-server operating according to the combined 'Join the Shortest Queue-Serve the Longest Queue' regime is analyzed. Both cases, with or without server's switch-over times, are investigated under the non-preemptive discipline. Instead of dealing with a state space comprised of two un-bounded dimensions, a non-conventional formulation is constructed, leading to an alternative two-dimensional state space, where only one dimension is infinite. As a result, the system is defined as a quasi birth and death process and is analyzed via both the probability generating functions method and the matrix geometric formulation. Consequently, the system's two-dimensional probability mass function is derived, from which the system's performance measures, such as mean queue sizes, mean sojourn times, fraction of time the server resides in each queue, correlation coefficient between the queue sizes, and the probability mass function of the difference between the queue sizes, are obtained. Extensive numerical results for various values of the system's parameters are presented, as well as a comparison between the current non-preemptive model and its twin system of preemptive service regime. One of the conclusions is that, depending on the variability of the various parameters, the preemptive regime is not necessarily more efficient than the non-preemptive one. Finally, economic issues are discussed and numerical comparisons are presented, showing the advantages and disadvantages of each regime.


Keywords Queueing • Join the Shortest Queue • Serve the Longest Queue • Polling • Probability generating functions • Matrix geometric

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## 1 Introduction

### 1.1 Background and contribution

The aim of the combined 'Join the Shortest Queue' (JSQ) and 'Serve the Longest Queue' (SLQ) service system is to equalize queue sizes. Each of the regimes, JSQ or SLQ, has been studied separately in the literature. Under the JSQ regime, the system is comprised of multiple separated queues, each with its specific service rate, while a newly arriving customer joins the shortest queue. In the SLQ model, a single server attends several queues, and always attempts to serve customers from the longest queue.

The combined JSQ-SLQ 2-queue Markovian system operating under the preemptive regime discipline with zero switch-over times has been recently introduced and analyzed in Perel et al. (2020). The investigation was further extended in Perel et al. (2022) for a 3-queue system.

The current paper broadens and generalizes the analysis of the 2-queue combined JSQ-SLQ system in three directions: (i) by analyzing the non-preemptive policy; (ii) by considering non-zero server's switch-over times, and (iii) by dealing with economic dichotomies to determine which and when one regime is preferable upon the other.

The non-preemptive policy implies that the server's switching decisions are made only upon service completions. Switching then occurs to a non-served queue only if its size is larger than the size of the queue attended by the server. Two versions are investigated: (i) non-preemptive with zero switch-over times, and (ii) non-preemptive with non-zero switch-over times. The JSQ-SLQ system with non-preemptive switching policy differs from the classical multi-class preemptive or non-preemptive priority models where class priority is fixed. In contrast, in the current model, the priority levels change dynamically, affected by changes in queue sizes.

An example of the combined JSQ-SLQ model, taken from the healthcare domain, is given in Perel et al. (2020), describing a medical clinic with a single operating physician and several treatment rooms, each with its dedicated medical assisting staff. A newly arriving patient is directed to the treatment room with the shortest queue. When the physician becomes available, s/he consistently visits the room having the longest queue. Service rates may differ between the treatment rooms.

Another example may be taken from the area of road transportation. Consider a road with two junctions apart from each other. There are two possible routes between the two junctions. At the first one, a traffic navigation application directs each arriving vehicle according to the JSQ policy. At the second junction, a traffic control mechanism gives extra 'green-light time' to the route with the longest queue.

A third example where the combined JSQ-SLQ model is applicable rises from the area of self-service cashiers in a supermarket. A paying customer joins the shortest queue, while a dedicated supermarket employee assists customers from the longest queue.
Main contributions A conventional formulation of the combined 2-queue JSQ-SLQ model leads to a two dimensional Markovian queueing system, where each dimension represents the queue size of the corresponding queue. Evidently, this 2-dimensional setting involves two un-bounded queues. For example, Flatto (1989), Avrachenkov
et al. (2014) and Adan et al. (2016) applied boundary value problem technique to solve such problems, while Bright and Taylor (1995) applied a truncation method. In contrast, by using an unorthodox approach, we are able to derive the equilibrium joint probability mass function of the queue lengths. The innovation in our analysis is that rather than defining an un-bounded 2 dimensional state space describing the queue sizes, the analysis is based on a novel formulation that transforms the statespace into two dimensions, one finite, the other infinite, thus enabling the use of probability generating functions (PGFs) technique combined with a Matrix Geometric analysis. A second contribution is the analysis of the non-preemptive case, while a third contribution is the introduction of server's switch-over times. A fourth contribution is an economic analysis to determine which and when one regime is preferable upon the other.

### 1.2 Related work

Fixed priority models, as well as polling systems, have been studied extensively in the queueing literature (see e.g. Conway et al. (2003), Takagi (1986), Kella and Yechiali (1988), Yechiali (1993), Boon et al. (2011), along with the extensive references therein). Browne and Yechiali (1989) investigated a polling-type system with server's dynamic switching rules. Recently, Perel and Yechiali (2017) and Jolles et al. (2018) investigated a single-server two-queue systems where the server's switching decisions are threshold-based, depending on the evolving queue sizes.

A system with two $M / G / 1$-type queues under the SLQ regime and non-preemptive discipline was analyzed by Cohen (1987). Flatto (1989) investigated the SLQ system with two identical queues and server's preemptive switching policy. A SLQ system with $N$ symmetric queues and non-preemptive policy was analyzed by van Houtum et al. (1997). A Markovian non-symmetric 2-queue system was studied by Knessl and Yao (2013) under heavy traffic regime. A wireless network with SLQ mechanism was studied by Maguluri et al. (2014), while Pedarsani and Walrand (2016) studied an SLQ model for a multi-class network.

Queueing systems operating under the JSQ policy have also been studied extensively in the literature. Winston (1977) studied a system with Poisson arrivals and a finite number of identical servers, each serving its own queue, and proved that the discounted number of jobs completing service by some time $t$ is maximized under the JSQ policy. This result was further extended by Hordijk and Koole (1990), who considered general arrival process, batch arrivals and finite buffers. Halfin (1985) derived bounds for the probability distribution of the number of customers in a JSQ system comprised of two identical servers. In Adan et al. (1991a), a JSQ system comprised of two non-symmetric servers was analyzed by using the compensation method (iterative approach), while a similar model with jockeying between the queues was studied in Adan et al. (1991b) via a matrix geometric approach. Furthermore, for the 2-queue JSQ system with Poisson arrival and two non-symmetric Exponential servers, Cohen (1998) derived explicitly the bivariate probability generating function of the stationary joint distribution of the queue lengths. This model was further analyzed in Adan et al. (2016), by using the compensation method and by solving a boundary value prob-
lem. van Houtum et al. (2001) studied a production system consisting of a group of parallel machines (servers) and multiple job types, where upon arrival, each job joins the shortest queue among all queues capable of serving it. Using numerical methods and truncation, the authors derive upper and lower bounds for the mean waiting time. Yao and Knessl $(2005,2006)$ studied a system comprised of two $M / M / \infty$ queues, in which new arrivals join the shortest queue. They applied asymptotic analysis for the bivariate generating function of the number of customers in each queue. A JSQ system operating under the Halfin-Whitt regime was studied by Eschenfeldt and Gamarnik (2018) and by Braverman (2020), where steady-state characteristics were obtained. A JSQ model with a large number of queues was studied in Dawson et al. (2019) and limiting results of various performance measures were derived. Dimitriou (2021) explored the stability and tail asymptotics of a Markovian single server retrial system with two infinite capacity orbits, where arriving customers that find the server busy join the shortest orbit queue. It should be emphasized that in the majority of the aforementioned JSQ models, a dedicated server is assigned to each queue. In deviation, in the current study, a unified JSQ-SLQ model is analyzed to examine a single-server polling-type system, where the server's transitions are from a shorter queue to a longer one, while incoming customers opt for the shortest queue.
Order of the paper and results This paper is organized as follows. Sections 2, 3 and 4 present and analyze the non-preemptive model with zero switch-over times. The system is first formulated in Sect. 2, while in Sect. 3 the partial probability generating functions of the system's states are derived, along with calculations of system's performance measures. In Sect.4, the system's stability condition is derived via the matrix geometric method. Section 5 analyzes the case with non-zero switch-over times. Extensive numerical results are presented in Sect. 6, and the two regimes - preemptive vs. non-preemptive - are compared. Various insights are then drawn, one of which is that the preemptive regime is not necessarily more efficient than the non-preemptive one. A detailed economic analysis follows, assessing the relative advantages of the two regimes and identifies the parameter values under which one regime performs better than the other. Section 7 concludes the paper.

## 2 Zero switch-over times

Consider a single server polling-type system comprised of two asymmetrical queues, denoted by $Q_{1}$ and $Q_{2}$. The arrival process of customers to the system is Poisson with rate $\lambda$. Service duration of an arbitrary customer in $Q_{i}$ is exponentially distributed with mean $1 / \mu_{i}, i=1,2$. An arriving customer follows the 'Join the Shortest Queue' (JSQ) policy, i.e. s/he always joins the shortest queue, while if both queues are with equal lengths, the customer joins $Q_{i}$ w.p. $p_{i} \geq 0$, where $p_{1}+p_{2}=1$. The server resides in each queue according to a non-preemptive 'Serve the Longest Queue' (SLQ) policy. That is, if service is rendered to one of the queues, and the number of customers in the un-served queue exceeds the number of customers in the served queue, the server first finishes the current service, and only then switches to the other (un-served) queue, provided that the number of customers in the un-served queue is still greater than the number of customers in the served queue. Otherwise, the server remains in the current
queue. Also, if the server has completed service in $Q_{i}(i=1,2)$ and both queues are equal in size, the server does not switch. In this case, if a new customer arrives before the next service completion, the customer will join $Q_{1}$ or $Q_{2}$ with probability $p_{1}$ and $p_{2}$, respectively, independently of the servers' position.

Let $L_{i}(t)$ denote the number of customers present in $Q_{i}(i=1,2)$ at time $t>0$, and let $I(t)$ denote the server's position at that time. $I(t)=i$ implies that the server resides in $Q_{i}$. Under stability (see stability condition in the sequel), let $L_{i}=\lim _{t \rightarrow \infty} L_{i}(t)$ and $I=\lim _{t \rightarrow \infty} I(t)$. Define $D(t)=\left(L_{1}(t)-L_{2}(t)\right)_{I(t)}$ and $D=\left(L_{1}-L_{2}\right)_{I}$. In deviation from the preemptive JSQ-SLQ regime, where $L_{1}-L_{2}$ may assume only the values $(-1),(0)$ and (1), under the non-preemptive combined JSQ-SLQ policy, $L_{1}-L_{2}$ assumes the values (-2), (-1), (0), (1) or (2). As a result from the model definition, $L_{1}-L_{2}=2$ necessarily implies that $I=1$, while $L_{1}-L_{2}=-2$ directly leads to $I=2$. However, when $L_{1}-L_{2}=-1,0,1$, the position of the server, $I$, may assume any of the values 1 or 2 , which leads to the form $1_{i}, 0_{i},-1_{i}$, for $i=1,2$. Hence, overall, there are 8 values for $D$, as follows:

- $D=2: L_{1}-L_{2}=2$, server is at $Q_{1}$, an arriving customer joins $Q_{2}$.
- $D=1_{i}: L_{1}-L_{2}=1$, server is at $Q_{i}(i=1,2)$, an arriving customer joins $Q_{2}$.
- $D=0_{i}: L_{1}=L_{2}$, server is at $Q_{i}(i=1,2)$, an arriving customer joins $Q_{i}$ w.p. $p_{i}\left(p_{1}+p_{2}=1\right)$.
- $D=-1_{i}: L_{1}-L_{2}=-1$, server is at $Q_{i}(i=1,2)$, an arriving customer joins $Q_{1}$.
- $D=-2: L_{1}-L_{2}=-2$, server is at $Q_{2}$, an arriving customer joins $Q_{1}$.

The above system is now formulated as a two dimensional continuous time Markovian process, having state space $\{(n, d)\}$, for $n \geq 0$ and $d \in \mathfrak{D}=$ $\left\{2,1_{1}, 1_{2}, 0_{1}, 0_{2},-1_{1},-1_{2},-2\right\}$, where $n$ denotes the number of customers in $Q_{1}$. In a stable system, the steady state joint probability mass function is denoted by $P_{n, d}=\mathbb{P}\left(L_{1}=n, D=d\right)$. A transition rate diagram of the process $\left(L_{1}(t), L_{2}(t)\right)$ is depicted in Fig. 1, from which the states of the resulting process $\left(L_{1}(t), D(t)\right)$ are readily concluded. The construction of Fig. 1 deserves a further explanation. It is comprised of squares along five diagonals. Each square represents a possible state of $\left(L_{1}, L_{2}\right)$. A square along the main diagonal has two possible states $(n, n)$ : one for the case where $D=0_{1}$ (both queues equal in size and the server resides in $Q_{1}$ ), and $D=0_{2}$ (the server resides in $Q_{2}$ ). Furthermore, each square on the diagonal above the main diagonal represents the cases where $D=-1_{i}$, where the index $i(i=1,2)$ indicates the position of the server. Finally, each square on the diagonal above the latter (the most north-east diagonal) represents the case where $D=-2$, while the diagonal below the main and the one below it represent the cases where $D=1_{i}$ and $D=2$, respectively.

## 3 Probability generating functions

This section is devoted to the derivation of the steady-state probability mass function of the 2 -dimensional process $\left(L_{1}, D\right)$. Following Perel et al. (2020), we use an unconventional construction of the probability generating functions (PGFs) and utilize


Fig. 1 Transition rate diagram of $\left(L_{1}(t), L_{2}(t)\right)$
their properties. Specifically, instead of deriving PGFs along the horizontal axis, as is commonly done, we derive PGFs along the diagonals of Fig. 1.

### 3.1 Steady-state equations and corresponding PGFs

For each diagonal of Fig. 1, we write the balance equations for all states $n$ along the diagonal. We first obtain the equations describing the states when the server resides
in $Q_{1}$, as follows: (i) When $d=2$ (lower diagonal),

$$
\begin{equation*}
\left(\lambda+\mu_{1}\right) P_{n, 2}=\mu_{2} P_{n, 1_{2}}, n \geq 2 \tag{1}
\end{equation*}
$$

(ii) When $d=1_{1}$ (states $1_{1}$ along the diagonal below the main),

$$
\begin{align*}
& \left(\lambda+\mu_{1}\right) P_{1,1_{1}}=\lambda p_{1}\left(P_{0,0_{1}}+P_{0,0_{2}}\right)+\mu_{1} P_{2,2}+\mu_{2} P_{1,0_{2}},  \tag{2}\\
& \left(\lambda+\mu_{1}\right) P_{n, 1_{1}}=\lambda p_{1} P_{n-1,0_{1}}+\lambda P_{n, 2}+\mu_{1} P_{n+1,2}+\mu_{2} P_{n, 0_{2}}, n \geq 2 . \tag{3}
\end{align*}
$$

(iii) For $d=0_{1}$ (states $0_{1}$ along the main diagonal),

$$
\begin{align*}
& \lambda P_{0,0_{1}}=\mu_{1} P_{1,1_{1}},  \tag{4}\\
& \left(\lambda+\mu_{1}\right) P_{1,0_{1}}=\lambda P_{1,1_{1}}+\mu_{1} P_{2,1_{1}},  \tag{5}\\
& \left(\lambda+\mu_{1}\right) P_{n, 0_{1}}=\lambda P_{n, 1_{1}}+\lambda P_{n-1,-1_{1}}+\mu_{1} P_{n+1,1_{1}}, n \geq 2 . \tag{6}
\end{align*}
$$

(iv) For $d=-1_{1}$ (states $-1_{1}$ along the diagonal above the main)

$$
\begin{equation*}
\left(\lambda+\mu_{1}\right) P_{n,-1_{1}}=\lambda p_{2} P_{n, 0_{1}}, n \geq 1 . \tag{7}
\end{equation*}
$$

In a similar manner, the balance equations for the case when the server resides in $Q_{2}$ are given by:
(v) When $d=1_{2}$ (states $1_{2}$ along the diagonal below the main),

$$
\begin{equation*}
\left(\lambda+\mu_{2}\right) P_{n, 1_{2}}=\lambda p_{1} P_{n-1,0_{2}}, n \geq 2 \tag{8}
\end{equation*}
$$

(vi) When $d=0_{2}$ (states $0_{2}$ along the main diagonal),

$$
\begin{align*}
& \lambda P_{0,0_{2}}=\mu_{2} P_{0,-1_{2}},  \tag{9}\\
& \left(\lambda+\mu_{2}\right) P_{1,0_{2}}=\lambda P_{0,-1_{2}}+\mu_{2} P_{1,-1_{2}},  \tag{10}\\
& \left(\lambda+\mu_{2}\right) P_{n, 0_{2}}=\lambda P_{n-1,-1_{2}}+\lambda P_{n, 1_{2}}+\mu_{2} P_{n,-1_{2}}, n \geq 2 . \tag{11}
\end{align*}
$$

(vii) For $d=-1_{2}$ (states $-1_{2}$ along the diagonal above the main),

$$
\begin{align*}
& \left(\lambda+\mu_{2}\right) P_{0,-1_{2}}=\lambda p_{2}\left(P_{0,0_{1}}+P_{0,0_{2}}\right)+\mu_{2} P_{0,-2}+\mu_{1} P_{1,0_{1}},  \tag{12}\\
& \left(\lambda+\mu_{2}\right) P_{n,-1_{2}}=\lambda p_{2} P_{n, 0_{2}}+\lambda P_{n-1,-2}+\mu_{2} P_{n,-2}+\mu_{1} P_{n+1,0_{1}}, n \geq 1 . \tag{13}
\end{align*}
$$

(viii) For $d=-2$ (upper diagonal),

$$
\begin{equation*}
\left(\lambda+\mu_{2}\right) P_{n,-2}=\mu_{1} P_{n+1,-1_{1}}, n \geq 0 \tag{14}
\end{equation*}
$$

For each $d \in \mathfrak{D}$, the conditional PGF of the number of customers in $Q_{1}$ is defined as follows:

$$
G_{d}(z)=\sum_{n} P_{n, d} z^{n}, \quad d \in \mathfrak{D},|z| \leq 1
$$

Multiplying each of the Eqs. (1)-(14) by $z^{n}$ (for the appropriate $n$ ), and summing over $n \geq 1$, we get a set of 8 equations for the probability generating functions, $\left(\left(G_{2}(z), G_{1_{1}}(z), G_{0_{1}}(z), G_{-1_{1}}(z), G_{1_{2}}(z), G_{0_{2}}(z), G_{-1_{2}}(z), G_{-2}(z)\right)\right.$, as follows:

$$
\begin{align*}
&\left(\lambda+\mu_{1}\right) G_{2}(z)= \mu_{2} G_{1_{2}}(z)  \tag{15}\\
&\left(\lambda+\mu_{1}\right) z G_{1_{1}}(z)=\left(\lambda z+\mu_{1}\right) G_{2}(z)+\lambda p_{1} z^{2} G_{0_{1}}(z)+\mu_{2} z G_{0_{2}}(z) \\
&-\left(\mu_{2}-\lambda p_{1}\right) z P_{0,0_{2}},  \tag{16}\\
&\left(\lambda+\mu_{1}\right) z G_{0_{1}}(z)=\left(\lambda z+\mu_{1}\right) G_{1_{1}}(z)+\lambda z^{2} G_{-1_{1}}(z)+\mu_{1} z P_{0,0_{1}},  \tag{17}\\
&\left(\lambda+\mu_{1}\right) G_{-1_{1}}(z)= \lambda p_{2} G_{0_{1}}(z)-\lambda p_{2} P_{0,0_{1}},  \tag{18}\\
&\left(\lambda+\mu_{2}\right) G_{1_{2}}(z)= \lambda p_{1} z G_{0_{2}}(z)-\lambda p_{1} P_{0,0_{2}},  \tag{19}\\
&\left(\lambda+\mu_{2}\right) G_{0_{2}}(z)=\lambda G_{1_{2}}(z)+\left(\lambda z+\mu_{2}\right) G_{-1_{2}}(z)+\mu_{2} P_{0,0_{2}},  \tag{20}\\
&\left(\lambda+\mu_{2}\right) z G_{-1_{2}}(z)= \mu_{1} G_{0_{1}}(z)+\lambda p_{2} z G_{0_{2}}(z)+\left(\lambda z+\mu_{2}\right) z G_{-2}(z) \\
&-\left(\mu_{1}-\lambda p_{2}\right) z P_{0,0_{1}},  \tag{21}\\
&\left(\lambda+\mu_{2}\right) z G_{-2}(z)= \mu_{1} G_{-1_{1}}(z) . \tag{22}
\end{align*}
$$

The set of 8 Eqs. (15)-(22) is condensed into a matrix form:

$$
\begin{equation*}
A(z) \cdot \vec{G}(z)=\vec{P}(z) \tag{23}
\end{equation*}
$$

where
$A(z)=\left(\begin{array}{cccccccc}\lambda+\mu_{1} & 0 & 0 & 0 & -\mu_{2} & 0 & 0 & 0 \\ -\left(\lambda z+\mu_{1}\right) & \left(\lambda+\mu_{1}\right) z & -\lambda p_{1} z^{2} & 0 & 0 & -\mu_{2} z & 0 & 0 \\ 0 & \left(-\lambda z+\mu_{1}\right) & \left(\lambda+\mu_{1}\right) z & -\lambda z^{2} & 0 & 0 & 0 & 0 \\ 0 & 0) & -\lambda p_{2} & \lambda+\mu_{1} & 0 & 0 & 0 & 0 \\ 0 & 0 & 0 & 0 & \lambda+\mu_{2} & -\lambda p_{1} z & 0 & 0 \\ 0 & 0 & 0 & 0 & -\lambda & \lambda+\mu_{2} & -\left(\lambda z+\mu_{2}\right) & 0 \\ 0 & 0 & -\mu_{1} & 0 & 0 & -\lambda p_{2} z & \left(\lambda+\mu_{2}\right) z & -\left(\lambda z+\mu_{2}\right) z \\ 0 & 0 & 0 & -\mu_{1} & 0 & 0 & 0 & \left(\lambda+\mu_{2}\right) z\end{array}\right)$,
$\vec{G}(z)=\left(G_{2}(z), G_{1_{1}}(z), G_{0_{1}}(z), G_{-1_{1}}(z), G_{1_{2}}(z), G_{0_{2}}(z), G_{-1_{2}}(z), G_{-2}(z)\right)^{T}$ is an 8 -dimensional column vector of the desired PGFs, and

$$
\begin{aligned}
\vec{P}(z)= & \left(\left(0,-\left(\mu_{2}-\lambda p_{1} z\right) z P_{0,0_{2}}, \mu_{1} z P_{0,0_{1}},-\lambda p_{2} P_{0,0_{1}},-\lambda p_{1} z P_{0,0_{2}}, \mu_{2} P_{0,0_{2}},\right.\right. \\
& \left.-\left(\mu_{1}-\lambda p_{2} z\right) P_{0,0_{1}}, 0\right)^{T}
\end{aligned}
$$

is a column vector containing the two unknown 'boundary probabilities', $P_{0,0_{1}}$ and $P_{0,0_{2}}$.

By Cramer's rule, $G_{d}(z)=\frac{\left|A_{d}(z)\right|}{\left|A^{\prime}(z)\right|}, d \in \mathfrak{D}=\left\{2,1_{1}, 1_{2}, 0_{1}, 0_{2},-1_{1},-1_{2},-2\right\}$, where $|A|$ is the determinant of a matrix $A$, and $A_{d}(z)$ is the matrix obtained from $A(z)$ by replacing the corresponding column of the latter matrix by $\vec{P}(z)$. Each of the PGFs $G_{d}(z), d \in \mathfrak{D}$, is a function of the probabilities $P_{0,0_{1}}$ and $P_{0,0_{2}}$, appearing in $\vec{P}(z)$. One equation for the calculation of the boundary probabilities, is the normalization
equation,

$$
\begin{equation*}
\sum_{d \in \mathfrak{D}} G_{d}(1)=\sum_{d \in \mathfrak{D}} \lim _{z \rightarrow 1} \frac{\left|A_{d}(z)\right|}{|A(z)|}=1, \tag{24}
\end{equation*}
$$

while a second relation between $P_{0,0_{1}}$ and $P_{0,0_{2}}$ is derived as follows. Since $G_{d}(z)$ is defined for all $|z| \leq 1$, each root of $|A(z)|$ is a root of $\left|A_{d}(z)\right|$. The determinant $|A(z)|$ is a 6 -degree polynomial which can be expressed as $|A(z)|=z^{2}(1-z) h(z)$, where

$$
\begin{equation*}
h(z)=-z^{3} \alpha_{3}+z^{2} \alpha_{2}-z \alpha_{1}-\alpha_{0} \tag{25}
\end{equation*}
$$

and

$$
\begin{aligned}
\alpha_{0}= & \mu_{1}^{2} \mu_{2}^{2}\left[\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2}+\lambda\left(\lambda+\mu_{1}\right)\left(\lambda+\mu_{2}\right)\left(\mu_{1} p_{1}+\mu_{2} p_{2}\right)+\lambda^{2} \mu_{1} \mu_{2} p_{1} p_{2}\right], \\
\alpha_{1}= & \lambda \mu_{1} \mu_{2}\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2}\left(\mu_{1}+\mu_{2}\right)+\mu_{1}^{2} \mu_{2}^{2}\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2} \\
& +\lambda^{2} \mu_{1}^{2} \mu_{2}^{2}\left(\lambda+\mu_{1}\right)\left(\lambda+\mu_{2}\right)+\lambda^{2} \mu_{1} \mu_{2}\left(\lambda+\mu_{1}\right)\left(\lambda+\mu_{2}\right)\left(\mu_{1} p_{1}+\mu_{2} p_{2}\right)\left(\mu_{1}+\mu_{2}\right) \\
& +\lambda \mu_{1}^{2} \mu_{2}^{2}\left(\lambda+\mu_{1}\right)\left(\lambda+\mu_{2}\right)\left(\mu_{1} p_{1}+\mu_{2} p_{2}\right)+\lambda^{2} \mu_{1}^{3} \mu_{2}^{3} p_{1} p_{2} \\
& +2 \lambda^{3} \mu_{1}^{2} \mu_{2}^{2} p_{1} p_{2}\left(\mu_{1}+\mu_{2}\right), \\
\alpha_{2}= & \lambda^{3}\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2}\left(\mu_{1}\left(1+p_{2}\right)+\mu_{2}\left(1+p_{1}\right)\right)+\lambda^{4}\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2} \\
& +\lambda^{2}\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2}\left(\mu_{1}^{2}+\mu_{2}^{2}\right)+\lambda^{4} \mu_{1} \mu_{2}\left(\lambda+\mu_{1}\right)\left(\lambda+\mu_{2}\right) \\
& +2 \lambda^{5} \mu_{1} \mu_{2} p_{1} p_{2}\left(\mu_{1}+\mu_{2}\right)+\lambda^{6} \mu_{1} \mu_{2} p_{1} p_{2}, \\
\alpha_{3}= & \lambda^{4}\left[\left(\lambda+\mu_{1}\right)^{2}\left(\lambda+\mu_{2}\right)^{2}-\lambda^{2} \mu_{1} \mu_{2} p_{1} p_{2}\right] .
\end{aligned}
$$

Note that the elements $\alpha_{i}, i=0, \ldots, 3$, are all positive and symmetric in $\mu_{1}$ and $\mu_{2}$ and in $p_{1}$ and $p_{2}$, as expected. The cubic polynomial $h(z)$ possesses 3 roots, denoted by $z_{k}, k=1,2,3$, that can be expressed explicitly by solving a cubic formula. Since $h(z)>0$ for all $z<-1$ and $h(0)<0$, then there is at least one root in the interval $(-1,0)$. By applying Descartes' rule of signs on $h(-z)$ (see e.g. Curtiss 1918), we conclude that $h(z)$ possesses a single root in $(-1,0)$, which we use as a second relation between the looked for boundary probabilities, $P_{0,0_{1}}$ and $P_{0,0_{2}}$. It can be verified that, under stability, the two other roots are greater than 1 , and therefore not relevant. Furthermore, the necessary and sufficient condition for the stability is is $h(1)<0$, and it is shown in Sect. 4.2 below, by using matrix geometric analysis. From all the above, $P_{0,0_{1}}$ and $P_{0,0_{2}}$ can be calculated, which provides us with expressions for the PGFs, $G_{d}(z)$, for all $d \in \mathfrak{D}$.

### 3.2 Performance measures

Define, respectively, the marginal probabilities of $D$ and of $L_{1}$ as

$$
\begin{aligned}
& P_{\bullet d}=\mathbb{P}(D=d)=\sum_{n=0}^{\infty} P_{n, d}=G_{d}(1), \quad d \in \mathfrak{D}, \\
& P_{n \bullet}=\mathbb{P}\left(L_{1}=n\right)=\sum_{d \in \mathfrak{D}} P_{n, d}, \quad n \geq 0 .
\end{aligned}
$$

Then,

$$
\begin{aligned}
& \mathbb{E}\left[L_{1}\right]=\sum_{n=0}^{\infty} n P_{n \bullet}=\sum_{d \in \mathfrak{D}} G_{d}^{\prime}(1), \\
& \begin{aligned}
\mathbb{E}[D] & =\mathbb{E}\left[L_{1}-L_{2}\right]=\sum_{d \in \mathfrak{D}} d G_{d}(1) \\
& =2 G_{2}(1)+G_{1_{1}}(1)+G_{1_{2}}(1)-G_{-1_{1}}(1)-G_{-1_{2}}(1)-2 G_{-2}(1) \\
\quad & =2\left(P_{\bullet 2}-P_{\bullet(-2)}\right)+\left(P_{\bullet 1_{1}}-P_{\bullet\left(-1_{1}\right)}\right)+\left(P_{\bullet 1_{2}}-P_{\bullet\left(-1_{2}\right)}\right),
\end{aligned}
\end{aligned}
$$

$$
\mathbb{E}\left[L_{2}\right]=\mathbb{E}\left[L_{1}\right]-\mathbb{E}[D]
$$

Furthermore,

$$
\begin{aligned}
\operatorname{Cov}\left(L_{1}, L_{2}\right) & =\mathbb{E}\left[L_{1} L_{2}\right]-\mathbb{E}\left[L_{1}\right] \mathbb{E}\left[L_{2}\right]=\mathbb{E}\left[L_{1}\left(L_{1}-D\right)\right]-\mathbb{E}\left[L_{1}\right] \mathbb{E}\left[L_{2}\right] \\
& =\mathbb{E}\left[L_{1}^{2}\right]-\mathbb{E}\left[L_{1} D\right]-\mathbb{E}\left[L_{1}\right] \mathbb{E}\left[L_{2}\right],
\end{aligned}
$$

where

$$
\begin{aligned}
\mathbb{E}\left[L_{1}^{2}\right]= & \sum_{d \in \mathfrak{D}} G_{d}^{\prime \prime}(1)+\mathbb{E}\left[L_{1}\right], \\
\mathbb{E}\left[L_{1} D\right]= & \sum_{d \in \mathfrak{D}} \sum_{n} n d P_{n, d}=2 \sum_{n} n P_{n, 2}-2 \sum_{n} n P_{n,-2}+\sum_{n} n P_{n, 1_{1}}-\sum_{n} n P_{n,-1_{1}} \\
& +\sum_{n} n P_{n, 1_{2}}-\sum_{n} n P_{n,-1_{2}} \\
= & 2\left[G_{2}^{\prime}(1)-G_{-2}^{\prime}(1)\right]+G_{1_{1}}^{\prime}(1)-G_{-1_{1}}^{\prime}(1)+G_{1_{2}}^{\prime}(1)-G_{-1_{2}}^{\prime}(1) .
\end{aligned}
$$

Also, the variance of $L_{i}$, for $i=1,2$, is given by

$$
\begin{aligned}
\operatorname{Var}\left(L_{1}\right) & =\mathbb{E}\left[L_{1}^{2}\right]-\left(\mathbb{E}\left[L_{1}\right]\right)^{2}, \\
\operatorname{Var}\left(L_{2}\right) & =\mathbb{E}\left[L_{2}^{2}\right]-\left(\mathbb{E}\left[L_{2}\right]\right)^{2}=\mathbb{E}\left[\left(L_{1}-D\right)^{2}\right]-\left(\mathbb{E}\left[L_{2}\right]\right)^{2} \\
& =\mathbb{E}\left[L_{1}^{2}\right]-2 \mathbb{E}\left[L_{1} D\right]+\mathbb{E}\left[D^{2}\right]-\left(\mathbb{E}\left[L_{2}\right]\right)^{2},
\end{aligned}
$$

where $\mathbb{E}\left[D^{2}\right]=\sum_{d \in \mathfrak{D}} d^{2} P_{\bullet} d=4\left(P_{\bullet 2}+P_{\bullet(-2)}\right)+P_{\bullet 1_{1}}+P_{\bullet\left(-1_{1}\right)}+P_{\bullet 1_{2}}+P_{\bullet\left(-1_{2}\right)}$.

As a result, the correlation coefficient between $L_{1}$ and $L_{2}$, denoted by $\operatorname{Cor}\left(L_{1}, L_{2}\right)$, can be explicitly calculated, using $\operatorname{Cor}\left(L_{1}, L_{2}\right)=\frac{\operatorname{Cov}\left(L_{1}, L_{2}\right)}{\sqrt{\operatorname{Var}\left(L_{1}\right) \operatorname{Var}\left(L_{2}\right)}}$ (for numerical results, see Sect. 6).

The effective arrival rate to $Q_{i}(i=1,2), \lambda_{e f f}^{i}$, is given by

$$
\begin{aligned}
& \lambda_{e f f}^{1}=\lambda\left(p_{1}\left(P_{\bullet 0_{1}}+P_{\bullet 0_{2}}\right)+P_{\bullet\left(-1_{1}\right)}+P_{\bullet\left(-1_{2}\right)}+P_{\bullet(-2)}\right), \\
& \lambda_{e f f}^{2}=\lambda\left(p_{2}\left(P_{\bullet 0_{1}}+P_{\bullet 0_{2}}\right)+P_{\bullet 1_{1}}+P_{\bullet 1_{2}}+P_{\bullet 2}\right) .
\end{aligned}
$$

Clearly, $\lambda_{e f f}^{1}+\lambda_{e f f}^{2}=\lambda$. The effective rate of work flowing into $Q_{i}$ is $\rho_{e f f}^{i}=\frac{\lambda_{e f f}^{i}}{\mu_{i}}$. We thus have,

$$
\mathbb{P}(\text { Server is idle })=P_{0,0_{1}}+P_{0,0_{2}}=1-\rho_{e f f}^{1}-\rho_{e f f}^{2} .
$$

By Little's Law, the mean sojourn time of a customer in $Q_{i}$, for $i=1,2$, is given by

$$
\mathbb{E}\left[W_{i}\right]=\frac{\mathbb{E}\left[L_{i}\right]}{\lambda_{\text {eff }}^{i}}
$$

## 4 Matrix geometric

### 4.1 Formulation

The system is formulated as a Quasi Birth and Death (QBD) process with 8 phases and an un-bounded number of levels, where phase $d$ corresponds to $D=d$, for $d \in \mathfrak{D}$, and each level $n$ corresponds to $L_{1}=n$, the total number of customers in $Q_{1}$. The connections between the PGF method and the matrix geometric approach have been investigated in several papers (see e.g. Perel and Yechiali 2013a, b; Paz and Yechiali 2014; Perel and Yechiali 2017; Phung-Duc 2017; Armony et al. 2019; Hanukov and Yechiali 2021).

For $n \geq 2$ define $\mathcal{S}_{n}$ to be the set of states
$\mathcal{S}_{n}=\left\{(n, 2),\left(n, 1_{1}\right),\left(n, 0_{1}\right),\left(n,-1_{1}\right),\left(n, 1_{2}\right),\left(n, 0_{2}\right),\left(n,-1_{2}\right),(n,-2)\right\}$, and arrange the system's states in the order

$$
\begin{aligned}
\mathcal{S}=\{ & \left(0,0_{1}\right),\left(0,0_{2}\right),\left(0,-1_{2}\right),(0,-2) ;\left(1,1_{1}\right),\left(1,0_{1}\right),\left(1,-1_{1}\right),\left(1,0_{2}\right),\left(1,-1_{2}\right), \\
& \left.(1,-2) ; \mathcal{S}_{2} ; \mathcal{S}_{3} ; \ldots ; \mathcal{S}_{n} \ldots\right\} .
\end{aligned}
$$

The infinitesimal generator matrix of the QBD , denoted by $Q$, is given by
where

$$
\begin{aligned}
& B_{1}^{0}=\left(\begin{array}{cccc}
-\lambda & 0 & \lambda p_{2} & 0 \\
0 & -\lambda & \lambda p_{2} & 0 \\
0 & \mu_{2} & -\left(\lambda+\mu_{2}\right) & 0 \\
0 & 0 & \mu_{2} & -\left(\lambda+\mu_{2}\right)
\end{array}\right), \quad B_{0}^{0}=\left(\begin{array}{cccccc}
\lambda p_{1} & 0 & 0 & 0 & 0 & 0 \\
\lambda p_{1} & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & \lambda & 0 & 0 \\
0 & 0 & 0 & 0 & \lambda & 0
\end{array}\right), \\
& B_{2}^{1}=\left(\begin{array}{cccc}
\mu_{1} & 0 & 0 & 0 \\
0 & 0 & \mu_{1} & 0 \\
0 & 0 & 0 & \mu_{1} \\
0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0
\end{array}\right), \\
& B_{1}^{1}=\left(\begin{array}{cccccc}
-\left(\lambda+\mu_{1}\right) & \lambda & 0 & 0 & 0 & 0 \\
0 & -\left(\lambda+\mu_{1}\right) & \lambda p_{2} & 0 & 0 & 0 \\
0 & 0 & -\left(\lambda+\mu_{1}\right) & 0 & 0 & 0 \\
\mu_{2} & 0 & 0 & -\left(\lambda+\mu_{2}\right) & \lambda p_{2} & 0 \\
0 & 0 & 0 & \mu_{2} & -\left(\lambda+\mu_{2}\right) & 0 \\
0 & 0 & 0 & 0 & \lambda & -\left(\lambda+\mu_{2}\right)
\end{array}\right), \\
& B_{0}^{1}=\left(\begin{array}{cccccccc}
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & \lambda & p_{1} & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & \lambda & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & \lambda p_{1} & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & \lambda & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & \lambda & 0
\end{array}\right), \quad B_{2}=\left(\begin{array}{cccccc}
\mu_{1} & 0 & 0 & 0 & 0 & 0 \\
0 & \mu_{1} & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & \mu_{1} & 0 \\
0 & 0 & 0 & 0 & 0 & \mu_{1} \\
0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0
\end{array}\right),
\end{aligned}
$$

and

$$
A_{1}=\left(\begin{array}{cccccccc}
-\left(\lambda+\mu_{1}\right) & \lambda & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & -\left(\lambda+\mu_{1}\right) & \lambda & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & -\left(\lambda+\mu_{1}\right) & \lambda p_{2} & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & -\left(\lambda+\mu_{1}\right) & 0 & 0 & 0 & 0 \\
\mu_{2} & 0 & 0 & 0 & -\left(\lambda+\mu_{2}\right) & \lambda & 0 & 0 \\
0 & \mu_{2} & 0 & 0 & 0 & -\left(\lambda+\mu_{2}\right) & \lambda p_{2} & 0 \\
0 & 0 & 0 & 0 & 0 & \mu_{2} & -\left(\lambda+\mu_{2}\right) & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & \mu_{2} & -\left(\lambda+\mu_{2}\right)
\end{array}\right)
$$

$$
A_{0}=\left(\begin{array}{cccccccc}
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & \lambda & p_{1} & 0 & 0 & 0 & 0 & 0 \\
0 \\
0 & 0 & \lambda & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & \lambda p_{1} & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & \lambda & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & \lambda & 0
\end{array}\right), \quad A_{2}=\left(\begin{array}{cccccccc}
0 & \mu_{1} & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & \mu_{1} & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & \mu_{1} & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & \mu_{1} \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & 0 & 0
\end{array}\right) .
$$

### 4.2 Stability condition

Define the matrix $A=A_{0}+A_{1}+A_{2}$. Then,

$$
A=\left(\begin{array}{cccccccc}
-\left(\lambda+\mu_{1}\right) & \lambda+\mu_{1} & 0 & 0 & 0 & 0 & 0 & 0 \\
0 & -\left(\lambda+\mu_{1}\right) & \lambda+\mu_{1} & 0 & 0 & 0 & 0 & 0 \\
0 & \lambda p_{1} & -\left(\lambda+\mu_{1}\right) & \lambda p_{2} & 0 & 0 & \mu_{1} & 0 \\
0 & 0 & \lambda & -\left(\lambda+\mu_{1}\right) & 0 & 0 & 0 & \mu_{1} \\
\mu_{2} & 0 & 0 & 0 & -\left(\lambda+\mu_{2}\right) & \lambda & 0 & 0 \\
0 & \mu_{2} & 0 & 0 & \lambda p_{1} & -\left(\lambda+\mu_{2}\right) & \lambda p_{2} & 0 \\
0 & 0 & 0 & 0 & 0 & \lambda+\mu_{2} & -\left(\lambda+\mu_{2}\right) & 0 \\
0 & 0 & 0 & 0 & 0 & 0 & \lambda+\mu_{2} & -\left(\lambda+\mu_{2}\right)
\end{array}\right),
$$

The matrix $A$ is the infinitesimal generator matrix of the process describing the evolution of $D$, given that $L_{1} \geq 2$. Let $\vec{\pi}$ be the stationary vector of the matrix $A$, i.e. $\vec{\pi} A=\overrightarrow{0}$ and $\vec{\pi} \cdot \vec{e}=1$ (where $\vec{e}$ is an 8 -dimensional column vector with all its entries equal to 1). From Neuts (1981), we have that the stability condition is

$$
\vec{\pi} A_{0} \vec{e}<\vec{\pi} A_{2} \vec{e}
$$

which, after tedious algebra, translates into a 5-degree polynomial in $\lambda$,

$$
\begin{align*}
g(\lambda)= & \lambda^{5}\left(\mu_{1}\left(p_{1}-2\right)+\mu_{2}\left(p_{2}-2\right)\right) \\
& +\lambda^{4}\left(p_{1} \mu_{1}^{2}+p_{2} \mu_{2}^{2}-3 \mu_{1}^{2}-3 \mu_{2}^{2}-4 \mu_{1} \mu_{2}-2 p_{1} p_{2} \mu_{1} \mu_{2}\right) \\
& -\lambda^{3}\left(\mu_{1}^{3}+\mu_{2}^{3}+\mu_{1}^{2} \mu_{2}+\mu_{1} \mu_{2}^{2}+p_{1} \mu_{1} \mu_{2}^{2}+p_{2} \mu_{1}^{2} \mu_{2}\right) \\
& +\lambda^{2}\left(p_{1} \mu_{1}^{3} \mu_{2}+p_{2} \mu_{1} \mu_{2}^{3}+7 \mu_{1}^{2} \mu_{2}^{2}-\left(p_{1}^{2}+p_{2}^{2}\right) \mu_{1}^{2} \mu_{2}^{2}\right) \\
& +\lambda\left(3 \mu_{1}^{3} \mu_{2}^{2}+3 \mu_{1}^{2} \mu_{2}^{3}+2 p_{1} \mu_{1}^{3} \mu_{2}^{2}+2 p_{2} \mu_{1}^{2} \mu_{2}^{3}\right)+2 \mu_{1}^{3} \mu_{2}^{3}>0 . \tag{26}
\end{align*}
$$

Note that Eq. (26) is symmetric in all of the system's parameters, i.e. $\lambda, \mu_{1}, \mu_{2}, p_{1}, p_{2}$. It can also be verified that Eq. (26) implies $h(1)<0$, where $h(z)$ is defined in Eq. (25). In order to explore the roots of $g(\lambda)$, we utilize again Descartes' rule of signs. Since the coefficients of $\lambda^{5}, \lambda^{4}$ and $\lambda^{3}$ are negative, while all other coefficients of $\lambda^{i}$, for $i=0,1,2$ are positive, there is a single change of signs between the coefficients. Therefore, there is a single positive root of this polynomial, which we denote by $\lambda_{0}=f\left(\mu_{1}, \mu_{2}, p_{1}, p_{2}\right)$. Since $g(0)=2 \mu_{1}^{3} \mu_{2}^{3}>0$, and $g(\infty)<0$, the system is
stable iff

$$
\begin{equation*}
\lambda<\lambda_{0} . \tag{27}
\end{equation*}
$$

In the symmetric case, where $\mu_{1}=\mu_{2}=\mu$, the stability condition (26) translates into

$$
\mu(\mu-\lambda)(\mu+\lambda)\left(3 \lambda^{3}+9 \lambda^{2} \mu+2 p_{1} p_{2} \lambda^{2} \mu+8 \lambda \mu^{2}+2 \mu^{3}\right)>0
$$

or, equivalently,

$$
\lambda<\mu .
$$

This holds for any value of $p_{1}$, since when both service rates are equal, the system can be looked upon as a single $M(\lambda) / M(\mu) / 1$ queue, for which the known stability condition is $\lambda<\mu$.

### 4.3 The equilibrium distribution

For $n \geq 0$ define the steady-state probability vector $\vec{P}_{n}$, as follows:

$$
\vec{P}_{n}= \begin{cases}\left(P_{0,0_{1}}, P_{0,0_{2}}, P_{0,-1_{2}}, P_{0,-2}\right), & n=0 \\ \left(P_{1,1_{1}}, P_{1,0_{1}}, P_{1,-1_{1}}, P_{1,0_{2}}, P_{1,-1_{2}}, P_{1,-2}\right), & n=1 \\ \left(P_{n, 2}, P_{n, 1_{1}}, P_{n, 0_{1}}, P_{n,-1_{1}}, P_{n, 1_{2}}, P_{n, 0_{2}}, P_{n,-1_{2}}, P_{n,-2}\right), & n \geq 2\end{cases}
$$

From Neuts (1981),

$$
\vec{P}_{n}=\vec{P}_{2} R^{n-2}, \quad n \geq 2
$$

where $R$ is the minimal non-negative solution of the matrix quadratic equation

$$
\begin{equation*}
A_{0}+R A_{1}+R^{2} A_{2}=0 \tag{28}
\end{equation*}
$$

The vectors $\vec{P}_{0}, \vec{P}_{1}$ and $\vec{P}_{2}$ are determined by the following linear system:

$$
\begin{align*}
& \vec{P}_{0} B_{1}^{0}+\vec{P}_{1} B_{2}^{1}=\overrightarrow{0} \\
& \vec{P}_{0} B_{0}^{0}+\vec{P}_{1} B_{1}^{1}+\vec{P}_{2} B_{2}=\overrightarrow{0} \\
& \vec{P}_{1} B_{0}^{1}+\vec{P}_{2} A_{1}+\vec{P}_{2} R A_{2}=\overrightarrow{0} \\
& \vec{P}_{0} \vec{e}_{0}+\vec{P}_{1} \vec{e}_{1}+\vec{P}_{2}[\mathbf{I}-R]^{-1} \vec{e}=1 \tag{29}
\end{align*}
$$

where $\vec{e}_{0}$ and $\vec{e}_{1}$ are 4-dimensional and 6-dimensional, respectively, vectors of 1 's. Equation (29) is the normalization equation.

The mean total number of customers in $Q_{1}, \mathbb{E}\left[L_{1}\right]$, is given by

$$
\begin{aligned}
\mathbb{E}\left[L_{1}\right] & =\sum_{n=1}^{\infty} n \vec{P}_{n} \vec{e}=\vec{P}_{1} \vec{e}_{1}+\sum_{n=2}^{\infty} n \vec{P}_{2} R^{n-2} \vec{e} \\
& =\vec{P}_{1} \vec{e}_{1}+\sum_{n=2}^{\infty}(n-1) \vec{P}_{2} R^{n-2} \vec{e}+\sum_{n=2}^{\infty} \vec{P}_{2} R^{n-2} \vec{e} \\
& =\vec{P}_{1} \vec{e}_{1}+\vec{P}_{2}\left([\mathbf{I}-R]^{-2}+[\mathbf{I}-R]^{-1}\right) \vec{e}
\end{aligned}
$$

### 4.4 Characterization of the rate matrix $R$

Equation (28) for the calculation of the matrix $R=\left[r_{i, j}\right]$ for $i, j=1, \ldots, 8$ involve 64 non-linear equations with 64 variables. Algorithms for calculatiog $R$ can be found in Neuts (1981), Latouche and Ramaswami (1999), Artalejo and Gómez-Corral (2008), and Harchol-Balter (2013). However, it is possible to characterize some of the properties of $R$. Following Ch. 6.2 in Latouche and Ramaswami (1999), the rate matrix $R$ can be represented as $R=A_{0} N$, where the element $N_{i j}$ of the matrix $N$ is the expected number of visits to state ( $n, j$ ), starting from state ( $n, i$ ), before the first visit to any of the states in levels lower than $n$. We recall that in our context, $L_{1}$ represents the levels, and the index $j$ refers to the phases (represented by $D$ ). Without calculating $N$, since the entries of the first, second and fifth rows of $A_{0}$ are all zeros, all elements in the corresponding rows of $R$ are zeros as well. That is, $r_{i, j}=0$ for $i=1,2,5$ and $j=1, \ldots, 8$. Furthermore, since the second and third rows of $A_{0}$ are equal, the second and third rows of $R$ will also be equal, namely $r_{2, j}=r_{3, j}, j=1,2,3,4$. In addition, from explicitly writing Eq. (28), each element of $R$ can be expressed in terms of only two elements, $r_{2,1}$ and $r_{2,2}$. These observations reduce the calculation efforts considerably. We also refer to Hanukov and Yechiali (2021), where it is shown that when the matrices $A_{0}, A_{1}$, and $A_{2}$ are all lower (or all upper) triangular, the elements of $R$ can be explicitly calculated and the stability condition is readily obtained.

## 5 Non-zero switch-over times

In this section we present the analysis for the 2-queue JSQ-SLQ system with non-zero switch-over times. Specifically, we assume that the server's switching time from $Q_{1}$ $\left(Q_{2}\right)$ to $Q_{2}\left(Q_{1}\right)$ is exponentially distributed with mean $1 / \gamma_{2}\left(1 / \gamma_{1}\right)$. We also assume that once the server switches from $Q_{i}$ to $Q_{j}$, it completes the switch and the service in $Q_{j}$, regardless the size of $Q_{i}$. After completing the service in $Q_{i}$, the server continues to operate according to the non-preemptive SLQ policy. The existence of switch-over times enlarges the number of possible states, such that the set $\mathfrak{D}$ consists now 16 states. Recall that the states in $\mathfrak{D}$ describe the difference $L_{1}-L_{2}$, as well as the location of the server, which, in the case with switch-over times, is not only a specific queue, but may also describe a transition phase of the server from $Q_{i}$ to $Q_{j}$. The various states are given as follows:

- $D=2$ : Server is at $Q_{1}$, an arriving customer joins $Q_{2}$.
- $D=2_{s_{1}}$ : Server switches to $Q_{1}$, an arriving customer joins $Q_{2}$.
- $D=1_{i}$ : Server is at $Q_{i}(i=1,2)$, an arriving customer joins $Q_{2}$.
- $D=1_{s_{i}}$ : Server switches to $Q_{i}(i=1,2)$, an arriving customer joins $Q_{2}$.
- $D=0_{i}$ : Server is at $Q_{i}(i=1,2)$, an arriving customer joins $Q_{i}$ w.p. $p_{i}\left(p_{1}+p_{2}=\right.$ 1).
- $D=0_{s_{i}}$ : Server switches to $Q_{i}(i=1,2)$, an arriving customer joins $Q_{i}$ w.p. $p_{i}$ $\left(p_{1}+p_{2}=1\right)$.
- $D=-1_{i}$ : Server is at $Q_{i}(i=1,2)$, an arriving customer joins $Q_{1}$.
- $D=-1_{s_{i}}$ : Server switches to $Q_{i}(i=1,2)$, an arriving customer joins $Q_{1}$.
- $D=-2$ : Server is at $Q_{2}$, an arriving customer joins $Q_{1}$.
- $D=-2_{s_{2}}$ : Server switches to $Q_{2}$, an arriving customer joins $Q_{1}$.

Similarly, as done in Sect.3, we write the balance equations of the corresponding system. The balance equations for the case when the server resides in $Q_{1}$ or switches to $Q_{1}$ are given by

When $d=2$,

$$
\begin{equation*}
\left(\lambda+\mu_{1}\right) P_{n, 2}=\gamma_{1} P_{n, 2_{s_{1}}}, n \geq 2 . \tag{30}
\end{equation*}
$$

When $d=2_{s_{1}}$,

$$
\begin{equation*}
\left(\lambda+\gamma_{1}\right) P_{n, 2_{s_{1}}}=\mu_{2} P_{n, 1_{2}}, n \geq 2 \tag{31}
\end{equation*}
$$

When $d=1_{1}$,

$$
\begin{align*}
& \left(\lambda+\mu_{1}\right) P_{1,1_{1}}=\lambda p_{1} P_{0,0_{1}}+\mu_{1} P_{2,2}+\gamma_{1} P_{1,1_{s_{1}}}  \tag{32}\\
& \left(\lambda+\mu_{1}\right) P_{n, 1_{1}}=\lambda p_{1} P_{n-1,0_{1}}+\lambda P_{n, 2}+\mu_{1} P_{n+1,2}+\gamma_{1} P_{n, 1_{s_{1}}}, n \geq 2 \tag{33}
\end{align*}
$$

When $d=1_{s_{1}}$,

$$
\begin{align*}
& \left(\lambda+\gamma_{1}\right) P_{1,1_{s_{1}}}=\lambda p_{1} P_{0,0_{2}}+\mu_{2} P_{1,0_{2}}  \tag{34}\\
& \left(\lambda+\gamma_{1}\right) P_{n, 1_{s_{1}}}=\lambda P_{n, 2_{s_{1}}}+\mu_{2} P_{n, 0_{2}}+\lambda p_{1} P_{n-1,0_{s_{1}}}, n \geq 2 \tag{35}
\end{align*}
$$

For $d=0_{1}$,

$$
\begin{align*}
\lambda P_{0,0_{1}} & =\mu_{1} P_{1,1_{1}}  \tag{36}\\
\left(\lambda+\mu_{1}\right) P_{1,0_{1}} & =\lambda P_{1,1_{1}}+\mu_{1} P_{2,1_{1}}+\gamma_{1} P_{1,0 s_{1}},  \tag{37}\\
\left(\lambda+\mu_{1}\right) P_{n, 0_{1}} & =\lambda P_{n, 1_{1}}+\lambda P_{n-1,-1_{1}}+\mu_{1} P_{n+1,1_{1}}+\gamma_{1} P_{n, 0 s_{1}}, n \geq 2 \tag{38}
\end{align*}
$$

For $d=0_{s_{1}}$,

$$
\begin{align*}
& \left(\lambda+\gamma_{1}\right) P_{1,0_{s_{1}}}=\lambda P_{1,1_{s_{1}}}  \tag{39}\\
& \left(\lambda+\gamma_{1}\right) P_{n, 0_{s_{1}}}=\lambda P_{n, 1_{s_{1}}}+\lambda P_{n-1,-1_{s_{1}}}, n \geq 2 \tag{40}
\end{align*}
$$

For $d=-1_{1}$,

$$
\begin{equation*}
\left(\lambda+\mu_{1}\right) P_{n,-1_{1}}=\lambda p_{2} P_{n, 0_{1}}+\gamma_{1} P_{n,-1_{s_{1}}}, n \geq 1 . \tag{41}
\end{equation*}
$$

For $d=-1_{s_{1}}$,

$$
\begin{equation*}
\left(\lambda+\gamma_{1}\right) P_{n,-1_{s_{1}}}=\lambda p_{2} P_{n, 0_{s_{1}}}, n \geq 1 . \tag{42}
\end{equation*}
$$

Furthermore, the balance equations for the case when the server resides in $Q_{2}$ or switches to $Q_{2}$ are given by:

When $d=1_{s_{2}}$,

$$
\begin{equation*}
\left(\lambda+\gamma_{2}\right) P_{n, 1_{s_{2}}}=\lambda p_{1} P_{n-1,0_{s_{2}}}, n \geq 2 . \tag{43}
\end{equation*}
$$

When $d=1_{2}$,

$$
\begin{equation*}
\left(\lambda+\mu_{2}\right) P_{n, 1_{2}}=\lambda p_{1} P_{n-1,0_{2}}+\gamma_{2} P_{n, 1_{s_{2}}}, n \geq 2 . \tag{44}
\end{equation*}
$$

When $d=0_{s_{2}}$,

$$
\begin{align*}
& \left(\lambda+\gamma_{2}\right) P_{1,0_{s_{2}}}=\lambda P_{0,-1_{s_{2}}}  \tag{45}\\
& \left(\lambda+\gamma_{2}\right) P_{n, 0_{s_{2}}}=\lambda P_{n-1,-1_{s_{2}}}+\lambda P_{n, 1_{s_{2}}}, n \geq 2 \tag{46}
\end{align*}
$$

When $d=0_{2}$,

$$
\begin{align*}
\lambda P_{0,0_{2}} & =\mu_{2} P_{0,-1_{2}}  \tag{47}\\
\left(\lambda+\mu_{2}\right) P_{1,0_{2}} & =\lambda P_{0,-1_{2}}+\mu_{2} P_{1,-1_{2}}+\gamma_{2} P_{1,0_{s_{2}}}  \tag{48}\\
\left(\lambda+\mu_{2}\right) P_{n, 0_{2}} & =\lambda P_{n-1,-1_{2}}+\lambda P_{n, 1_{2}}+\mu_{2} P_{n,-1_{2}}+\gamma_{2} P_{n, 0_{s_{2}}}, n \geq 2 \tag{49}
\end{align*}
$$

For $d=-1_{s_{2}}$,

$$
\begin{align*}
& \left(\lambda+\gamma_{2}\right) P_{0,-1_{s_{2}}}=\lambda p_{2} P_{0,0_{1}}+\mu_{1} P_{1,0_{1}}  \tag{50}\\
& \left(\lambda+\gamma_{2}\right) P_{n,-1_{s_{2}}}=\lambda p_{2} P_{n, 0_{s_{2}}}+\lambda P_{n-1,-2_{s_{2}}}+\mu_{1} P_{n+1,0_{1}}, n \geq 1 \tag{51}
\end{align*}
$$

For $d=-1_{2}$,

$$
\begin{align*}
& \left(\lambda+\mu_{2}\right) P_{0,-1_{2}}=\lambda p_{2} P_{0,0_{2}}+\mu_{2} P_{0,-2}+\gamma_{2} P_{0,-1_{s_{2}}}  \tag{52}\\
& \left(\lambda+\mu_{2}\right) P_{n,-1_{2}}=\lambda p_{2} P_{n, 0_{2}}+\lambda P_{n-1,-2}+\mu_{2} P_{n,-2}+\gamma_{2} P_{n,-1_{s_{2}}}, n \geq 1 \tag{53}
\end{align*}
$$

For $d=-2_{s_{2}}$,

$$
\begin{equation*}
\left(\lambda+\gamma_{2}\right) P_{n,-2_{s_{2}}}=\mu_{1} P_{n+1,-1_{1}}, n \geq 0 . \tag{54}
\end{equation*}
$$

For $d=-2$,

$$
\begin{equation*}
\left(\lambda+\mu_{2}\right) P_{n,-2}=\gamma_{2} P_{n,-2_{s_{2}}}, n \geq 0 . \tag{55}
\end{equation*}
$$

The set of 16 PGFs is derived as follows:

$$
\begin{align*}
& \left(\lambda+\mu_{1}\right) G_{2}(z)=\gamma_{1} G_{2_{s_{1}}}(z),  \tag{56}\\
& \left(\lambda+\gamma_{1}\right) G_{2_{s_{1}}}(z)=\mu_{2} G_{1_{2}}(z),  \tag{57}\\
& \left(\lambda+\mu_{1}\right) z G_{1_{1}}(z)=\lambda p_{1} z^{2} G_{0_{1}}(z)+\left(\lambda z+\mu_{1}\right) G_{2}(z)+\gamma_{1} z G_{1_{s_{1}}}(z)  \tag{58}\\
& \left(\lambda+\gamma_{1}\right) G_{1_{s_{1}}}(z)=\lambda G_{2_{s_{1}}}(z)+\mu_{2} G_{0_{2}}(z)+\lambda p_{1} z G_{0_{s_{1}}}(z)-\left(\mu_{2}-\lambda p_{1} z\right) P_{0,0_{2}}  \tag{59}\\
& \left(\lambda+\mu_{1}\right) z G_{0_{1}}(z)=\left(\lambda z+\mu_{1}\right) G_{1_{1}}(z)+\lambda z^{2} G_{-1_{1}}(z)+\gamma_{1} z G_{0_{s_{1}}}(z)+\mu_{1} z P_{0,0_{1}},  \tag{60}\\
& \left(\lambda+\gamma_{1}\right) G_{0_{s_{1}}}(z)=\lambda G_{1_{s_{1}}}(z)+\lambda z G_{-1_{s_{1}}}(z),  \tag{61}\\
& \left(\lambda+\mu_{1}\right) G_{-1_{1}}(z)=\lambda p_{2} G_{0_{1}}(z)+\gamma_{1} G_{-1_{s_{1}}}(z)-\lambda p_{2} P_{0,0_{1}},  \tag{62}\\
& \left(\lambda+\gamma_{1}\right) G_{-1_{s_{1}}}(z)=\lambda p_{2} G_{0_{s_{1}}}(z),  \tag{63}\\
& \left(\lambda+\gamma_{2}\right) G_{1_{s_{2}}}(z)=\lambda p_{1} z G_{0_{s_{2}}}(z),  \tag{64}\\
& \left(\lambda+\mu_{2}\right) G_{1_{2}}(z)=\lambda p_{1} z G_{0_{2}}(z)+\gamma_{2} G_{1_{s_{2}}}(z)-\lambda p_{1} z P_{0,0_{2}},  \tag{65}\\
& \left(\lambda+\gamma_{2}\right) G_{0_{s_{2}}}(z)=\lambda z G_{-1_{s_{2}}}(z)+\lambda G_{1_{s_{2}}}(z),  \tag{66}\\
& \left(\lambda+\mu_{2}\right) G_{0_{2}}(z)=\lambda G_{1_{2}}(z)+\left(\lambda z+\mu_{2}\right) G_{-1_{2}}(z)+\gamma_{2} G_{0_{s_{2}}}(z)+\mu_{2} P_{0,0_{2}},  \tag{67}\\
& \left(\lambda+\gamma_{2}\right) z G_{-1_{s_{2}}}(z)=\mu_{1} G_{0_{1}}(z)+\lambda p_{2} z G_{0_{s_{2}}}(z)+\lambda z^{2} G_{-2_{s_{2}}}(z)-\left(\mu_{1}-\lambda p_{2} z\right) P_{0,0_{1}}, \tag{68}
\end{align*}
$$

$$
\begin{equation*}
\left(\lambda+\mu_{2}\right) G_{-1_{2}}(z)=\lambda p_{2} G_{0_{2}}(z)+\left(\lambda z+\mu_{2}\right) G_{-2}(z)+\gamma_{2} G_{-1_{s_{2}}}(z), \tag{69}
\end{equation*}
$$

$$
\begin{equation*}
\left(\lambda+\gamma_{2}\right) z G_{-2_{s_{2}}}(z)=\mu_{1} G_{-1_{1}}(z), \tag{70}
\end{equation*}
$$

$$
\begin{equation*}
\left(\lambda+\mu_{2}\right) G_{-2}(z)=\gamma_{2} G_{-2_{s_{2}}}(z) \tag{71}
\end{equation*}
$$

Note that in order to fully obtain the above PGFs that satisfy Eqs. (56)-(71), only two probabilities, $P_{0,0_{1}}$ and $P_{0,0_{2}}$, need to ba calculated. The derivation of these probabilities is done similarly to the way described in Sect. 3.1. Furthermore, all performance measures calculated in Sect. 3.2 for the case without switch-over times can be derived in a similar manner for the model with switch-over times. Therefore, we omit these calculations from the paper. Nevertheless, in Sect. 6 we present numerical results for various performance measures of this model.

## 6 Numerical results

In this section we compare the performance measures of the current non-preemptive JSQ-SLQ model with the corresponding measures of the preemptive JSQ-SLQ queue-
ing system studied in Perel et al. (2020). For each model, we present numerical results of the system's performance measures for a wide range of parameters. Note that all calculations are based on the analytical results and derived in the paper. There are two sets of tables: Tables 1, 2, 3, 4 and 5, comparing various performance measures, and Tables $6,7,8,9$ and 10 , dealing with the probability distribution of $D$. Furthermore, numerical results for the case with server's switch-over times are provided in Tables 12 and 13.

### 6.1 Comparison between preemptive and non-preemptive JSQ-SLQ models without switch-over times

Tables 1, 2, 3, 4 and 5 present sets of results, where the calculated measures in all tables are $\mathbb{E}\left[L_{i}\right], \mathbb{E}\left[W_{i}\right], \lambda_{\text {eff }}^{i}, \rho_{\text {eff }}^{i}(i=1,2)$ and $\operatorname{Cor}\left(L_{1}, L_{2}\right)$. The tables maintain the same parameter values: $\lambda=4$ and $\mu_{2}=5$, but differ by the values of the parameter $p_{1}$, where $p_{1}=0.2,0.5,0.8,1$ in Tables $1,2,3$ and 4, respectively, while $p_{1}=\frac{\mu_{1}}{\mu_{1}+\mu_{2}}$ in Table 5. In each table, the values of $\mu_{1}$ vary between 3.3 and 100 . For each value of $\mu_{1}$, the first row describes results for the preemptive model, while the second row presents results for the non-preemptive case. Furthermore, for each case, we present the stability condition in terms of $\lambda$, according to the values of $\mu_{1}, \mu_{2}, p_{1}$ and $p_{2}$ (see Eq. 27).

### 6.1.1 Insights from Tables 1, 2, 3, 4 and 5

1. In all tables, when $\mu_{2}=5, \mathbb{E}\left[L_{1}\right]$ and $\mathbb{E}\left[L_{2}\right]$, as well as $\mathbb{E}\left[W_{1}\right]$ and $\mathbb{E}\left[W_{2}\right]$, are all decreasing functions of $\mu_{1}$.
2. When $p_{1} \leq 0.5$ (Tables 1 and 2 ) and for small values of $\mu_{1}$, the values for $\mathbb{E}\left[L_{i}\right]$ and $\mathbb{E}\left[W_{i}\right](i=1,2)$ in the non-preemptive model are significantly larger than the corresponding measures in the preemptive case. However, as $\mu_{1}$ increases, the differences between $\mathbb{E}\left[L_{i}\right], i=1,2$, and between $\mathbb{E}\left[W_{i}\right]$ in both models become negligible.
3. When $p_{1}>0.5$ (Tables 3 and 4) and for small values of $\mu_{1}$, the results for $\mathbb{E}\left[L_{i}\right]$ and for $\mathbb{E}\left[W_{i}\right]$ in the non-preemptive model are smaller than the corresponding results in the preemptive case. This can be explained since whenever $p_{1}>0.5$, in case when $L_{1}=L_{2}$, arriving customers are more likely to join $Q_{1}$ rather than $Q_{2}$. In the preemptive case, when $\mu_{1}$ is relatively small (large mean service time), the server switches to $Q_{2}$ immediately when $L_{2}>L_{1}$, even before completing service at $Q_{1}$. While the server resides at $Q_{2}$, customers arrive at $Q_{1}$, causing $L_{1}$ to exceed $L_{2}$, so the server immediately switches back to $Q_{1}$, and so on. Note that $\mu_{2}$ is not relatively large, therefore, the server rapidly alternates between the queues, significant amount of times, without completing service. On the other hand, in the non-preemptive case, the server first completes service, and only then, if needed, switches to the other queue. Intuitively, one may think that switching over in the middle of a service of a job (under a non-resume policy) may always be worse than the non-preemptive case. However, the above results show that this is not always the case, and, depending on the parameter values, switching may prove beneficial.
Table 1 Numerical results for $\lambda=4, \mu_{2}=5, p_{1}=0.2, p_{2}=0.8$

| $\mu_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{e f f}^{1}$ | $\lambda_{e f f}^{2}$ | $\rho_{e f f}^{1}$ | $\rho_{e f f}^{2}$ | $\operatorname{Cor}\left(L_{1}, L_{2}\right)$ | stab. cond |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3.3 | 9.49 | 9.53 | 6.52 | 3.75 | 1.46 | 2.54 | 0.44 | 0.51 | 0.9975 | $\lambda<4.2$ |
|  | 22.52 | 22.67 | 13.14 | 9.92 | 1.71 | 2.29 | 0.52 | 0.46 | 0.9990 | $\lambda<4.08$ |
| 3.5 | 6.23 | 6.28 | 4.24 | 2.48 | 1.47 | 2.53 | 0.42 | 0.51 | 0.9943 | $\lambda<4.31$ |
|  | 9.22 | 9.38 | 5.42 | 4.08 | 1.70 | 2.30 | 0.49 | 0.46 | 0.9945 | $\lambda<4.21$ |
| 4 | 3.46 | 3.53 | 2.31 | 1.41 | 1.50 | 2.50 | 0.37 | 0.50 | 0.9836 | $\lambda<4.56$ |
|  | 3.85 | 4.00 | 2.29 | 1.72 | 1.68 | 2.32 | 0.42 | 0.46 | 0.9715 | $\lambda<4.50$ |
| 4.5 | 2.46 | 2.55 | 1.62 | 1.03 | 1.52 | 2.48 | 0.34 | 0.49 | 0.9708 | $\lambda<4.79$ |
|  | 2.52 | 2.67 | 1.51 | 1.14 | 1.66 | 2.34 | 0.37 | 0.47 | 0.9403 | $\lambda<4.76$ |
| 5 | 1.95 | 2.05 | 1.27 | 0.83 | 1.54 | 2.46 | 0.31 | 0.49 | 0.9575 | $\lambda<5$ |
|  | 1.92 | 2.08 | 1.16 | 0.88 | 1.65 | 2.35 | 0.33 | 0.47 | 0.9084 | $\lambda<5$ |
| 8 | 0.98 | 1.14 | 0.61 | 0.48 | 1.60 | 2.40 | 0.20 | 0.48 | 0.8927 | $\lambda<5.98$ |
|  | 0.95 | 1.12 | 0.58 | 0.48 | 1.63 | 2.37 | 0.20 | 0.47 | 0.7926 | $\lambda<6.04$ |
| 20 | 0.51 | 0.73 | 0.3 | 0.32 | 1.67 | 2.33 | 0.08 | 0.46 | 0.8153 | $\lambda<8$ |
|  | 0.54 | 0.75 | 0.33 | 0.32 | 1.65 | 2.35 | 0.08 | 0.47 | 0.7381 | $\lambda<7.93$ |
| 100 | 0.34 | 0.61 | 0.20 | 0.27 | 1.71 | 2.29 | 0.02 | 0.45 | 0.8040 | $\lambda<10.88$ |
|  | 0.40 | 0.65 | 0.24 | 0.28 | 1.67 | 2.33 | 0.02 | 0.47 | 0.7512 | $\lambda<10.01$ |

Table 2 Numerical results for $\lambda=4, \mu_{2}=5, p_{1}=p_{2}=0.5$

| $\mu_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{\text {eff }}^{1}$ | $\lambda_{e f f}^{2}$ | $\rho_{e f f}^{1}$ | $\rho_{e f f}^{2}$ | $\operatorname{Cor}\left(L_{1}, L_{2}\right)$ | stab. cond |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3.3 | 61.46 | 61.39 | 32.99 | 28.72 | 1.87 | 2.13 | 0.56 | 0.43 | 0.9999 | $\lambda<4.03$ |
|  | 6013.3 | 6013.3 | 3099.1 | 2919.6 | 1.94 | 2.06 | 0.59 | 0.41 | 1.00 | $\lambda<4.00$ |
| 3.5 | 12.52 | 12.46 | 6.64 | 5.89 | 1.88 | 2.12 | 0.53 | 0.42 | 0.9985 | $\lambda<4.16$ |
|  | 14.90 | 14.88 | 7.65 | 7.25 | 1.95 | 2.05 | 0.56 | 0.41 | 0.9977 | $\lambda<4.13$ |
| 4 | 4.35 | 4.31 | 2.25 | 2.08 | 1.93 | 2.07 | 0.48 | 0.41 | 0.9888 | $\lambda<4.46$ |
|  | 4.45 | 4.48 | 2.26 | 2.18 | 1.97 | 2.03 | 0.49 | 0.41 | 0.9763 | $\lambda<4.45$ |
| 4.5 | 2.71 | 2.69 | 1.37 | 1.32 | 1.97 | 2.03 | 0.44 | 0.41 | 0.9738 | $\lambda<4.74$ |
|  | 2.71 | 2.71 | 1.37 | 1.35 | 1.98 | 2.02 | 0.44 | 0.40 | 0.9421 | $\lambda<4.73$ |
| 5 | 2 | 2 | 1 | 1 | 2 | 2 | 0.4 | 0.4 | 0.9565 | $\lambda<5$ |
|  | 2 | 2 | 1 | 1 | 2 | 2 | 0.4 | 0.4 | 0.9042 | $\lambda<5$ |
| 8 | 0.87 | 0.92 | 0.41 | 0.49 | 2.12 | 1.88 | 0.26 | 0.38 | 0.8568 | $\lambda<6.27$ |
|  | 0.91 | 0.94 | 0.44 | 0.49 | 2.07 | 1.93 | 0.26 | 0.39 | 0.7493 | $\lambda<6.20$ |
| 20 | 0.37 | 0.51 | 0.16 | 0.29 | 2.28 | 1.72 | 0.11 | 0.35 | 0.7130 | $\lambda<9.21$ |
|  | 0.45 | 0.55 | 0.21 | 0.30 | 2.19 | 1.81 | 0.11 | 0.36 | 0.6596 | $\lambda<8.47$ |
| 100 | 0.20 | 0.39 | 0.08 | 0.24 | 2.38 | 1.62 | 0.02 | 0.32 | 0.6973 | $\lambda<14.85$ |
|  | 0.29 | 0.44 | 0.13 | 0.26 | 2.30 | 1.70 | 0.02 | 0.34 | 0.6698 | $\lambda<11.12$ |

Table 3 Numerical results for $\lambda=4, \mu_{2}=5, p_{1}=0.8, p_{2}=0.2$

| $\mu_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{\text {eff }}^{1}$ | $\lambda_{\text {eff }}^{2}$ | $\rho_{\text {eff }}^{1}$ | $\rho_{\text {eff }}^{2}$ | $\operatorname{Cor}\left(L_{1}, L_{2}\right)$ | stab. cond |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3.5 | 130.99 | 130.83 | 57.23 | 76.46 | 2.29 | 1.71 | 0.65 | 0.34 | 0.9999 | $\lambda<4.01$ |
|  | 35.19 | 34.98 | 16.18 | 19.16 | 2.17 | 1.83 | 0.62 | 0.37 | 0.9996 | $\lambda<4.05$ |
| 4 | 5.65 | 5.51 | 2.40 | 3.36 | 1.64 | 2.36 | 0.59 | 0.33 | 0.9933 | $\lambda<4.36$ |
|  | 5.17 | 4.99 | 2.31 | 2.83 | 2.24 | 1.76 | 0.56 | 0.35 | 0.9820 | $\lambda<4.40$ |
| 4.5 | 2.97 | 2.86 | 1.23 | 1.80 | 1.59 | 2.41 | 0.54 | 0.32 | 0.9778 | $\lambda<4.69$ |
|  | 2.92 | 2.74 | 1.27 | 1.61 | 2.30 | 1.70 | 0.51 | 0.34 | 0.9473 | $\lambda<4.71$ |
| 5 | 2.05 | 1.95 | 0.83 | 1.27 | 2.46 | 1.54 | 0.49 | 0.31 | 0.9575 | $\lambda<5$ |
|  | 2.08 | 1.92 | 0.87 | 1.16 | 2.34 | 1.66 | 0.47 | 0.33 | 0.9056 | $\lambda<5$ |
| 8 | 0.78 | 0.73 | 0.29 | 0.54 | 2.66 | 1.34 | 0.33 | 0.27 | 0.8215 | $\lambda<6.57$ |
|  | 0.87 | 0.77 | 0.34 | 0.53 | 2.54 | 1.46 | 0.32 | 0.29 | 0.7074 | $\lambda<6.35$ |
| 20 | 0.28 | 0.30 | 0.09 | 0.29 | 2.94 | 1.06 | 0.15 | 0.21 | 0.5803 | $\lambda<10.70$ |
|  | 0.36 | 0.34 | 0.13 | 0.30 | 2.85 | 1.15 | 0.14 | 0.23 | 0.5473 | $\lambda<9.04$ |
| 100 | 0.09 | 0.17 | 0.03 | 0.22 | 3.19 | 0.81 | 0.03 | 0.16 | 0.5326 | $\lambda<21.82$ |
|  | 0.17 | 0.21 | 0.05 | 0.24 | 3.12 | 0.88 | 0.03 | 0.18 | 0.5190 | $\lambda<12.27$ |
|  |  |  |  |  |  |  |  |  |  |  |

Table 4 Numerical results for $\lambda=4, \mu_{2}=5, p_{1}=1, p_{2}=0$

| $\mu_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{e f f}^{1}$ | $\lambda_{e f f}^{2}$ | $\rho_{e f f}^{1}$ | $\rho_{e f f}^{2}$ | $\operatorname{Cor}\left(L_{1}, L_{2}\right)$ | stab. cond |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 4.2 | 4.68 | 4.48 | 1.75 | 3.36 | 2.67 | 1.33 | 0.63 | 0.27 | 0.9907 | $\lambda<4.44$ |
|  | 4.23 | 3.93 | 1.72 | 2.55 | 2.45 | 1.55 | 0.58 | 0.31 | 0.9744 | $\lambda<4.98$ |
| 4.5 | 3.17 | 2.98 | 1.17 | 2.31 | 2.71 | 1.29 | 0.60 | 0.26 | 0.9807 | $\lambda<4.66$ |
|  | 3.06 | 2.77 | 1.22 | 1.86 | 2.51 | 1.49 | 0.57 | 0.30 | 0.9526 | $\lambda<4.69$ |
| 5 | 2.08 | 1.91 | 0.75 | 1.56 | 2.77 | 1.23 | 0.55 | 0.25 | 0.9593 | $\lambda<5$ |
|  | 2.13 | 1.87 | 0.83 | 1.31 | 2.58 | 1.42 | 0.52 | 0.28 | 0.9097 | $\lambda<5$ |
| 8 | 0.73 | 0.61 | 0.24 | 0.63 | 3.03 | 0.97 | 0.38 | 0.19 | 0.8027 | $\lambda<6.79$ |
|  | 0.84 | 0.65 | 0.29 | 0.59 | 2.89 | 1.11 | 0.36 | 0.22 | 0.6783 | $\lambda<6.47$ |
| 20 | 0.24 | 0.17 | 0.07 | 0.31 | 3.44 | 0.56 | 0.17 | 0.11 | 0.4857 | $\lambda<11.86$ |
|  | 0.30 | 0.20 | 0.09 | 0.32 | 3.37 | 0.63 | 0.17 | 0.13 | 0.4085 | $\lambda<9.44$ |
| 100 | 0.05 | 0.03 | 0.01 | 0.22 | 3.85 | 0.15 | 0.039 | 0.029 | 0.2878 | $\lambda<29.22$ |
|  | 0.06 | 0.04 | 0.02 | 0.24 | 3.83 | 0.17 | 0.04 | 0.03 | 0.2056 | $\lambda<13.07$ |

Table 5 Numerical results for $\lambda=4, \mu_{2}=5, p_{1}=\frac{\mu_{1}}{\mu_{1}+\mu_{2}}, p_{2}=\frac{\mu_{2}}{\mu_{1}+\mu_{2}}$

| $\mu_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{e f f}^{1}$ | $\lambda_{e f f}^{2}$ | $\rho_{\text {eff }}^{1}$ | $\rho_{e f f}^{2}$ | $\operatorname{Cor}\left(L_{1}, L_{2}\right)$ | stab. cond |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 3.3 | 22.00 | 21.97 | 12.75 | 9.65 | 1.72 | 2.27 | 0.52 | 0.45 | 0.9995 | $\lambda<4.09$ |
|  | 67.08 | 67.11 | 35.94 | 31.45 | 1.87 | 2.13 | 0.56 | 0.42 | 0.9998 | $\lambda<4.03$ |
| 3.5 | 9.75 | 9.72 | 5.53 | 4.35 | 1.76 | 2.24 | 0.50 | 0.45 | 0.9976 | $\lambda<4.20$ |
|  | 12.69 | 12.71 | 6.75 | 5.99 | 1.87 | 2.13 | 0.54 | 0.42 | 0.9968 | $\lambda<4.16$ |
| 4 | 4.16 | 4.14 | 2.25 | 1.93 | 1.85 | 2.15 | 0.46 | 0.43 | 0.9878 | $\lambda<4.48$ |
|  | 4.34 | 4.35 | 2.26 | 2.09 | 1.91 | 2.09 | 0.48 | 0.42 | 0.9753 | $\lambda<4.46$ |
| 4.5 | 2.69 | 2.67 | 1.39 | 1.29 | 1.93 | 2.07 | 0.43 | 0.41 | 0.9735 | $\lambda<4.74$ |
|  | 2.69 | 2.70 | 1.38 | 1.32 | 1.96 | 2.04 | 0.43 | 0.41 | 0.9417 | $\lambda<4.74$ |
| 5 | 2 | 2 | 1 | 1 | 2 | 2 | 0.4 | 0.4 | 0.9565 | $\lambda<5$ |
|  | 2 | 2 | 1 | 1 | 2 | 2 | 0.4 | 0.4 | 0.9042 | $\lambda<5$ |
| 8 | 0.82 | 0.85 | 0.35 | 0.51 | 2.22 | 1.67 | 0.29 | 0.33 | 0.8428 | $\lambda<6.38$ |
|  | 0.89 | 0.88 | 0.40 | 0.50 | 2.25 | 1.75 | 0.28 | 0.35 | 0.7332 | $\lambda<6.26$ |
| 20 | 0.28 | 0.30 | 0.09 | 0.29 | 2.94 | 1.06 | 0.15 | 0.21 | 0.5803 | $\lambda<10.70$ |
|  | 0.36 | 0.35 | 0.13 | 0.30 | 2.85 | 1.15 | 0.14 | 0.23 | 0.5457 | $\lambda<9.04$ |
| 100 | 0.060 | 0.067 | 0.01 | 0.21 | 3.68 | 0.32 | 0.03 | 0.06 | 0.3762 | $\lambda<27.21$ |
|  | 0.09 | 0.08 | 0.02 | 0.24 | 3.65 | 0.35 | 0.03 | 0.07 | 0.3295 | $\lambda<12.88$ |

Table 6 The probability mass function of $D$, for $\lambda=4, \mu_{2}=5, p_{1}=0.2, p_{2}=0.8$

| $\mu_{1}$ | 2 | $1_{1}$ | $0_{1}$ | $-1_{1}$ | $1_{2}$ | $0_{2}$ | $-1_{2}$ | -2 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3.3 | 0.0123 | 0.1782 | 0.2367 | 0.1002 | 0.0179 | 0.2172 | 0.2007 | 0.0368 |
| 3.5 | 0.0118 | 0.1732 | 0.2301 | 0.0901 | 0.0177 | 0.2348 | 0.2073 | 0.035 |
| 4 | 0.0108 | 0.1620 | 0.2181 | 0.0706 | 0.0173 | 0.2694 | 0.2204 | 0.0314 |
| 5 | 0.0093 | 0.1438 | 0.2062 | 0.0467 | 0.0168 | 0.3141 | 0.2372 | 0.0259 |
| 8 | 0.0067 | 0.1083 | 0.2023 | 0.0189 | 0.0161 | 0.3722 | 0.2587 | 0.0168 |
| 20 | 0.0032 | 0.055 | 0.2261 | 0.0029 | 0.0154 | 0.4164 | 0.2747 | 0.0063 |
| 100 | 0.0007 | 0.0129 | 0.26 | 0.0001 | 0.0151 | 0.431 | 0.2792 | 0.001 |

Table 7 The probability mass function of $D$, for $\lambda=4, \mu_{2}=5, p_{1}=p_{2}=0.5$

| $\mu_{1}$ | 2 | $1_{1}$ | $0_{1}$ | $-1_{1}$ | $1_{2}$ | $0_{2}$ | $-1_{2}$ | -2 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3.3 | 0.0276 | 0.2243 | 0.2639 | 0.0723 | 0.0403 | 0.1815 | 0.1636 | 0.0265 |
| 3.5 | 0.0263 | 0.219 | 0.2611 | 0.0655 | 0.0395 | 0.1952 | 0.1679 | 0.0255 |
| 4 | 0.0236 | 0.2071 | 0.2574 | 0.0522 | 0.0376 | 0.2226 | 0.1763 | 0.0232 |
| 5 | 0.0195 | 0.1871 | 0.2582 | 0.0352 | 0.0352 | 0.2582 | 0.1871 | 0.0195 |
| 8 | 0.0131 | 0.1458 | 0.2804 | 0.0143 | 0.0316 | 0.3027 | 0.1994 | 0.0127 |
| 20 | 0.0058 | 0.0782 | 0.3526 | 0.0019 | 0.0280 | 0.3264 | 0.2027 | 0.0044 |
| 100 | 0.0012 | 0.0192 | 0.4347 | 0.00005 | 0.026 | 0.3219 | 0.1964 | 0.00055 |

Table 8 The probability mass function of $D$, for $\lambda=4, \mu_{2}=5, p_{1}=0.8, p_{2}=0.2$

| $\mu_{1}$ | 2 | $1_{1}$ | $0_{1}$ | $-1_{1}$ | $1_{2}$ | $0_{2}$ | $-1_{2}$ | -2 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3.5 | 0.0379 | 0.2707 | 0.2904 | 0.0301 | 0.0568 | 0.1654 | 0.137 | 0.0117 |
| 4 | 0.0329 | 0.2583 | 0.2968 | 0.0244 | 0.0526 | 0.1836 | 0.1405 | 0.0109 |
| 5 | 0.0259 | 0.2372 | 0.3141 | 0.0168 | 0.0467 | 0.2062 | 0.1438 | 0.0093 |
| 8 | 0.0156 | 0.1913 | 0.3728 | 0.0069 | 0.0374 | 0.2271 | 0.1427 | 0.0062 |
| 20 | 0.0057 | 0.1089 | 0.5192 | 0.0009 | 0.0272 | 0.2116 | 0.1245 | 0.002 |
| 100 | 0.001 | 0.0282 | 0.6821 | 0.00001 | 0.0203 | 0.16999 | 0.0983 | 0.0002 |

Table 9 The probability mass function of $D$, for $\lambda=4, \mu_{2}=5, p_{1}=1, p_{2}=0$

| $\mu_{1}$ | 2 | $1_{1}$ | $0_{1}$ | $-1_{1}$ | $1_{2}$ | $0_{2}$ | $-1_{2}$ | -2 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 4.2 | 0.0354 | 0.2924 | 0.3296 | 0.00 | 0.0581 | 0.1646 | 0.1199 | 0.00 |
| 4.5 | 0.0324 | 0.2858 | 0.3386 | 0.00 | 0.0552 | 0.1686 | 0.1194 | 0.00 |
| 5 | 0.0283 | 0.2754 | 0.3541 | 0.00 | 0.051 | 0.1731 | 0.1181 | 0.00 |
| 8 | 0.015 | 0.2272 | 0.4439 | 0.00 | 0.036 | 0.1723 | 0.1056 | 0.00 |
| 20 | 0.0038 | 0.1349 | 0.6602 | 0.00 | 0.0183 | 0.1164 | 0.0664 | 0.00 |
| 100 | 0.0002 | 0.0365 | 0.9083 | 0.00 | 0.0046 | 0.0323 | 0.0181 | 0.00 |

Table 10 The probability mass function of $D$, for $\lambda=4, \mu_{2}=5, p_{1}=\frac{\mu_{1}}{\mu_{1}+\mu_{2}}, p_{2}=\frac{\mu_{2}}{\mu_{1}+\mu_{2}}$

| $\mu_{1}$ | 2 | $1_{1}$ | $0_{1}$ | $-1_{1}$ | $1_{2}$ | $0_{2}$ | $-1_{2}$ | -2 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3.3 | 0.0227 | 0.208 | 0.255 | 0.0831 | 0.0332 | 0.1924 | 0.1751 | 0.0305 |
| 3.5 | 0.0224 | 0.205 | 0.2522 | 0.0739 | 0.0336 | 0.2057 | 0.1785 | 0.0287 |
| 4 | 0.0214 | 0.1983 | 0.2502 | 0.0563 | 0.0343 | 0.2306 | 0.1838 | 0.0251 |
| 5 | 0.0195 | 0.1871 | 0.2582 | 0.0352 | 0.0352 | 0.2582 | 0.1871 | 0.0195 |
| 8 | 0.0145 | 0.1623 | 0.3141 | 0.0118 | 0.0349 | 0.2745 | 0.1774 | 0.0105 |
| 20 | 0.0057 | 0.1089 | 0.5192 | 0.0009 | 0.0272 | 0.2116 | 0.1245 | 0.002 |
| 100 | 0.0005 | 0.0343 | 0.8487 | 0.00 | 0.0094 | 0.06846 | 0.0386 | 0.00004 |

4. As $\mu_{1}$ increases, the maximum value of $\lambda$ ensuring stability increases. However, for $p_{1} \leq 0.5$, the stability condition is more strict in the non-preemptive case than it is in the preemptive one.
5. For $\mu_{1} \leq 5$, as $p_{1}$ increases, $\mathbb{E}\left[L_{1}\right]$ and $\mathbb{E}\left[L_{2}\right]$ increase, as higher proportion of the arrivals joins $Q_{1}$ when $L_{1}=L_{2}$. If the server renders service in $Q_{1}$, the nonpreemptive policy forces the server to remain there until service completion, and if $\mu_{1}<\mu_{2}$, the number of customers arriving to the system during this service time, increases. However, when $\mu_{1}>\mu_{2}$, as $p_{1}$ increases, the values for $\mathbb{E}\left[L_{1}\right]$ and $\mathbb{E}\left[L_{2}\right]$ decrease.
6. The correlation coefficient between $L_{1}$ and $L_{2}$ is always positive. This follows from the JSQ policy. Furthermore, as $\mu_{1}$ increases, the correlation coefficient between $L_{1}$ and $L_{2}$ decreases. To explain this phenomenon, consider first small values of $\mu_{1}$, where the server is busy in $Q_{1}$ for a long period of time. Therefore, since customers follow the JSQ policy, the number of customers in both queues increases simultaneously. However, for large values of $\mu_{1}$, the behavior of $L_{2}$ is less affected by $L_{1}$, since the server resides in $Q_{1}$ a short amount of time.
7. Table 5 exhibits a balancing result between $\mathbb{E}\left[L_{1}\right]$ and $\mathbb{E}\left[L_{2}\right]$, in both the preemptive and the non-preemptive cases. When the joining probabilities are relative to the service rates ratio, i.e $p_{i}=\frac{\mu_{i}}{\mu_{1}+\mu_{2}}, i=1,2$, both mean queue lengths are significantly reduced.
8. All tables show that, when $\mu_{1}$ increases, the proportion of time the server spends in both queues (i.e. $\rho_{e f f}^{1}+\rho_{e f f}^{2}$ ) reduces significantly.
9. In the symmetric case, i.e. when $p_{1}=p_{2}=0.5$ and $\mu_{1}=\mu_{2}$, the obtained results correspond to an $M / M / 1$ system, as expected (see Tables 2 and 5 for the case $\mu_{1}=5$ ).

Figures 2 and 3 depict the impact of $p_{1}$ on the sum $\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]$, for the case $\lambda=4$, and $\mu_{2}=5$, where in Fig. $2 \mu_{1}=3.5$ while in Fig. $3 \mu_{1}=6.5$. Obviously, when $\mu_{1}<\mu_{2}$, that is, service time is longer in $Q_{1}$ than in $Q_{2}$ (Fig. 2), $\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]$ grows as $p_{1}$ increases. This occurs since when $p_{1}$ is high, customers tend to join the slower queue when both queue sizes are equal, which increases the total number of customers in the system. On the other hand, when $\mu_{1}>\mu_{2}$ (Fig. 3), $\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]$ decreases as $p_{1}$ increases, since joining the faster queue with high probability when the queue sizes are equal, reduces the total number of customers in the system. Moreover,


Fig. 2 The impact of $p_{1}$ on $\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]$ for $\lambda=4, \mu_{1}=3.5$ and $\mu_{2}=5$


Fig. 3 The impact of $p_{1}$ on $\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]$ for $\lambda=4, \mu_{1}=6.5$ and $\mu_{2}=5$
in Fig. $2\left(\mu_{1}<\mu_{2}\right)$, when $p_{1}<0.6$, the sum of queue sizes is almost equal in both preemptive and non-preemptive regimes, while when $p_{1}$ increases beyond 0.6 , the total queue under the preemptive regime becomes much higher than that under the non-preemptive regime. In Fig. $3\left(\mu_{1}>\mu_{2}\right)$, there is no much difference between the sum of the queue sizes as $p_{1}$ increases, but for lower values of $p_{1}$ the non-preemptive case is better than the preemptive regime and for $p_{1}$ larger, the opposite holds.

### 6.1.2 Insights from Tables 6, 7, 8, 9 and 10

Tables $6,7,8,9$ and 10 deal with the distribution of $D$, where $\lambda=4, \mu_{2}=5$ while $p_{1}$ and $\mu_{1}$ change.

1. In all tables, as $\mu_{1}$ increases, the probabilities $\mathbb{P}(D=2), \mathbb{P}(D=-2)$ and $\mathbb{P}(D=$ $1_{1}$ ) decrease drastically. Indeed, as service rate in $Q_{1}$ increases, the total number
of customers in both queues decreases, as well as the gap between the queue sizes. Furthermore, $\mathbb{P}\left(D=1_{2}\right)$ also decreases when $\mu_{1}$ increases, where in Tables 6,7 and 8 the decrease is more moderate than in Tables 9 and 10.
2. In all tables, the probability $\mathbb{P}\left(L_{1}=L_{2}\right)=\mathbb{P}\left(D=0_{1}\right)+\mathbb{P}\left(D=0_{2}\right)$ increases as $\mu_{1}$ increases, since a fast service rate enables the system to be more balanced.
3. When $p_{1} \geq p_{2}$ (Tables $\left.7,8,9,10\right)$, the probability $\mathbb{P}\left(D=0_{1}\right)$ increases as $\mu_{1}$ increases. Indeed, when service rate in $Q_{1}$ is high, as well as the proportion of customers joining $Q_{1}$ when $L_{1}=L_{2}$, the server spends more time in $Q_{1}$ and queue sizes tend to be more balanced. However, when $p_{1}<p_{2}$ as in Table 6, $\mathbb{P}\left(D=0_{1}\right)$ has no drastic changes for different values of $\mu_{1}$, while $\mathbb{P}\left(D=0_{2}\right)$ increases significantly as $\mu_{1}$ becomes large.
4. For the symmetric case (Tables 7 and 10 for $\mu_{1}=5$ ) we get, as expected, that $\mathbb{P}\left(\right.$ server at $\left.Q_{1}\right)=\mathbb{P}\left(\right.$ server at $\left.Q_{2}\right)=0.5$, where

$$
\mathbb{P}\left(\text { server at } Q_{1}\right)=\mathbb{P}(D=2)+\mathbb{P}\left(D=1_{1}\right)+\mathbb{P}\left(D=0_{1}\right)+\mathbb{P}\left(D=-1_{1}\right),
$$

$$
\mathbb{P}\left(\text { server at } Q_{2}\right)=\mathbb{P}\left(D=1_{2}\right)+\mathbb{P}\left(D=0_{2}\right)+\mathbb{P}\left(D=-1_{2}\right)+\mathbb{P}(D=-2) .
$$

5. In Table 9 , where $p_{1}=1$, we have that $\mathbb{P}(D=-2)=0$, since in case when $L_{1}=L_{2}$, arriving customers will always join $Q_{1}$, so $L_{2}$ can never exceed $L_{1}$ by more than one customer.

### 6.2 Economic comparison between the preemptive and non-preemptive regimes

Let $C$ denote the total operational cost rate of the JSQ-SLQ queueing systems with zero-switch-over times (either under the preemptive policy or under the nonpreemptive policy). Specifically, $C$ is the sum of the expected total cost per unit time incurred by customers' sojourn times in the system, and the cost per unit time resulting from the server's switch-overs. Define $\mathbb{E}\left[S_{i j}\right]$ to be the expected number of switches per unit time from $Q_{i}$ to $Q_{j}, i \neq j$, and let $\mathbb{E}[$ Switch $]=\mathbb{E}\left[S_{12}\right]+\mathbb{E}\left[S_{21}\right]$ denote the mean total number of server's switches per unit time. In steady state, we have $\mathbb{E}\left[S_{12}\right]=\mathbb{E}\left[S_{21}\right]$. Then, the expected total operational cost per unit time is

$$
\mathbb{E}[C]=c\left(\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]\right)+s \mathbb{E}[\text { Switch }],
$$

where $c$ is the cost rate for a customer's sojourn in the system, and $s$ is the cost per server's switch. Without loss of generality, we set $c=1$ and write

$$
\begin{equation*}
\mathbb{E}[C]=\left(\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]\right)+s \mathbb{E}[\text { Switch }] \tag{72}
\end{equation*}
$$

The measures $\mathbb{E}\left[S_{i j}\right]$ for $i \neq j$ are calculated as follows.
For the preemptive regime,

$$
\mathbb{E}\left[S_{12}\right]=\lambda p_{2} G_{0_{1}}(1)+\mu_{1}\left(G_{0_{1}}(1)-P_{0,0_{1}}\right),
$$

Table 11 Values of $\mathbb{E}[$ Switch $]$ for the preemptive and the non-preemptive regimes, where $\lambda=4$ and $\mu_{2}=5$

| $\mu_{1}$ | 3.5 | 4 | 4.5 | 5 | 5.5 | 6 | 6.5 | 7 | 7.5 | 8 | 8.5 | 9 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| $p_{1}=0.2$ | 2.89 | 2.83 | 2.77 | 2.73 | 2.7 | 2.67 | 2.65 | 2.93 | 2.61 | 2.59 | 2.58 | 2.57 |
|  | 2.23 | 2.24 | 2.25 | 2.26 | 2.26 | 2.27 | 2.27 | 2.27 | 2.27 | 2.28 | 2.28 | 2.28 |
| $p_{1}=0.5$ | 2.98 | 2.94 | 2.91 | 2.89 | 2.87 | 2.84 | 2.83 | 2.81 | 2.79 | 2.78 | 2.77 | 2.75 |
|  | 2.24 | 2.28 | 2.31 | 2.33 | 2.35 | 2.36 | 2.37 | 2.37 | 2.38 | 2.38 | 2.38 | 2.38 |
| $p_{1}=0.8$ | 2.81 | 2.75 | 2.76 | 2.73 | 2.71 | 2.68 | 2.65 | 2.62 | 2.59 | 2.56 | 2.54 | 2.51 |
|  | 2.2 | 2.23 | 2.25 | 2.26 | 2.26 | 2.25 | 2.24 | 2.23 | 2.22 | 2.2 | 2.19 | 2.18 |

$$
\mathbb{E}\left[S_{21}\right]=\lambda p_{1} G_{0_{2}}(1)+\mu_{2}\left(G_{0_{2}}(1)-P_{0,0_{2}}\right),
$$

while for the non-preemptive regime,

$$
\begin{aligned}
& \mathbb{E}\left[S_{12}\right]=\lambda p_{2} P_{0,0_{1}}+\mu_{1}\left(G_{0_{1}}(1)+G_{-1_{1}}(1)-P_{0,0_{1}}\right), \\
& \mathbb{E}\left[S_{21}\right]=\lambda p_{1} P_{0,0_{2}}+\mu_{2}\left(G_{0_{2}}(1)+G_{1_{2}}(1)-P_{0,0_{1}}\right) .
\end{aligned}
$$

### 6.2.1 Switching rates - numerical results and insights

The values of $\mathbb{E}[$ Switch $]$, for both the preemptive and non-preemptive regimes, for $\lambda=4$ and $\mu_{2}=5$, where $p_{1}$ assumes the values $0.2,0.5$ and 0.8 and $\mu_{1}$ varies between 3.5 to 10 , are given in Table 11. For each value of $p_{1}$, the top row represents the preemptive case while the bottom row refers to the non-preemptive case.

Insights from Table 11 It is evident from the table that the switching rate in the non-preemptive regime is smaller than the corresponding one in the preemptive case. This is a direct result of the non-preemptive policy. However, as $\mu_{1}$ increases, the gaps between the switch-over rates between the two regimes decrease. Furthermore, Table 11 shows that for the preemptive case, the switching rate is a decreasing function of $\mu_{1}$, while for the non-preemptive case, it moderately increases when $p_{1}=0.2$ and $p_{1}=0.5$, and concave when $p_{1}=0.8$.

### 6.2.2 Expected total operational cost-results and insights

Figures 4,5 and 6 depict $\mathbb{E}[C]$ (Eq. 72) as a function of $\mu_{1}$, both for the preemptive and the non-preemptive regimes, where in each figure $\lambda=4, \mu_{2}=5$ while $p_{1}$ assumes the values $0.2,0.5$ and 0.8 , respectively. Furthermore, for each figure, the left sub-figures describe the case where switch-over cost is $s=0.5$, the mid sub-figures refer to the case where $s=2$, while the right sub-figures describe the case where switch-over cost is $s=10$.

Insights from Figs. 4, 5 and 6:

1. In all 3 figures, the systems' operational cost is a decreasing function of $\mu_{1}$.
2. Figure 4 a and b show that when $p_{1}=0.2$, for small values of $\mu_{1}$, the cost rate in the preemptive case is much lower than in the non-preemptive case, since there are significant gaps between the number of customers in the system, i.e. between $\mathbb{E}\left[L_{1}\right]+\mathbb{E}\left[L_{2}\right]$, as shown in Table 1 . However, as $\mu_{1}$ increases, the graphs of


Fig. 4 Expected operational costs of the preemptive and non-preemptive regimes, for $\lambda=4, \mu_{2}=5$ and $p_{1}=0.2$


Fig. 5 Expected operational costs of the preemptive and non-preemptive regimes, for $\lambda=4, \mu_{2}=5$ and $p_{1}=0.5$


Fig. 6 Expected operational costs of the preemptive and non-preemptive regimes, for $\lambda=4, \mu_{2}=5$ and $p_{1}=0.8$
the costs of the two regimes interlace, and the operational cost rate of the nonpreemptive case becomes lower than the preemptive one. Furthermore, when the penalty on a server's switch is $s=10$ (Fig.4c), the differences between the costs are greater, but with smaller values for the non-preemptive case. The above findings result from the values of $\mathbb{E}[$ switch $]$, given in Table 11 .
3. When $p_{1}=0.5$ and $s=0.5$ (Fig. 5a), the graphs of operational cost rates of both regimes seem to be quite identical. As the penalty for a switch rises, see Fig. 5b and c , the expected operational cost of the non-preemptive regime becomes lower than the preemptive one.
4. Figure $6 \mathrm{a}-\mathrm{c}$ refer to the case where $p_{1}=0.8$, i.e., that when $L_{1}=L_{2}$, a newly arriving customer joins $Q_{1}$ with high probability. As a result, the operational cost rate of the non-preemptive system is smaller than the corresponding cost in preemptive case, especially when $s=10$.

### 6.3 Numerical results for non-zero switch-over times

Tables 12 and 13 present sets of results, where the calculated measures in both tables are again $\mathbb{E}\left[L_{i}\right], \mathbb{E}\left[W_{i}\right], \lambda_{e f f}^{i}, \rho_{e f f}^{i}(i=1,2)$ and $\operatorname{Cor}\left(L_{1}, L_{2}\right)$, as well as the measure $\mathbb{P}$ (Switch), which is the fraction of time that the server is switching from $Q_{i}$ to $Q_{j}$. In both tables, the parameters' values are: $\mu_{2}=5, p_{1}=p_{2}=0.5, \gamma_{2}=3$, while $\lambda_{1}=2$ in Table 12 and $\lambda_{1}=3$ in Table 13. Furthermore, in Table 12 the values of $\mu_{1}$ and $\gamma_{1}$ vary between 3 to 8 , and 3 to 10 , respectively, while in Table 13 the values of $\mu_{1}$ and $\gamma_{1}$ vary between 5 to 10 , and 10 to 14 , respectively.

## Insights from Tables 12 and 13:

1. In both tables, $\mathbb{E}\left[L_{1}\right]$ and $\mathbb{E}\left[L_{2}\right]$, as well as $\mathbb{E}\left[W_{1}\right]$ and $\mathbb{E}\left[W_{2}\right]$, are all decreasing functions of $\mu_{1}$ and $\gamma_{1}$.
Table 12 Numerical results for the model with switch-over times, $\lambda=2, \mu_{2}=5, p_{1}=p_{2}=0.5, \gamma_{2}=3$

| $\mu_{1}$ | $\gamma_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{\text {eff }}^{1}$ | $\lambda_{\text {eff }}^{2}$ | $\rho_{\text {eff }}^{1}$ | $\rho_{\text {eff }}^{2}$ | Cor $\left(L_{1}, L_{2}\right)$ | $\mathbb{P}($ Switch $)$ |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 3 | 3 | 6.40 | 6.38 | 6.54 | 6.24 | 0.98 | 1.02 | 0.37 | 0.20 | 0.9912 | 0.41 |
|  | 4 | 3.454 | 3.447 | 3.49 | 3.41 | 0.99 | 1.01 | 0.33 | 0.20 | 0.9727 | 0.36 |
|  | 5 | 2.66 | 2.66 | 2.67 | 2.65 | 0.998 | 1.002 | 0.33 | 0.20 | 0.9573 | 0.33 |
|  | 10 | 1.79 | 1.81 | 1.76 | 1.83 | 1.02 | 0.98 | 0.338 | 0.197 | 0.9205 | 0.27 |
| 4 | 3 | 2.72 | 2.71 | 2.74 | 2.68 | 0.99 | 1.01 | 0.25 | 0.20 | 0.9584 | 0.41 |
|  | 4 | 1.88 | 1.88 | 1.87 | 1.89 | 0.997 | 1.003 | 0.25 | 0.20 | 0.9236 | 0.36 |
|  | 5 | 1.56 | 1.57 | 1.54 | 1.59 | 1.01 | 0.99 | 0.253 | 0.198 | 0.8991 | 0.33 |
|  | 10 | 1.13 | 1.57 | 1.17 | 1.10 | 1.03 | 0.97 | 0.26 | 0.19 | 0.8479 | 0.27 |
| 5 | 3 | 1.95 | 1.95 | 1.95 | 1.95 | 1 | 1 | 0.20 | 0.20 | 0.9291 | 0.41 |
|  | 4 | 1.42 | 1.43 | 1.40 | 1.46 | 1.01 | 0.99 | 0.203 | 0.197 | 0.8853 | 0.36 |
|  | 5 | 1.21 | 1.23 | 1.18 | 1.26 | 1.02 | 0.98 | 0.205 | 0.195 | 0.8563 | 0.33 |
|  | 10 | 0.90 | 0.94 | 0.86 | 0.96 | 1.05 | 0.95 | 0.209 | 0.191 | 0.7993 | 0.27 |
| 6 | 3 | 1.62 | 1.63 | 1.61 | 1.64 | 1.007 | 0.993 | 0.18 | 0.20 | 0.9070 | 0.41 |
|  | 4 | 1.21 | 1.23 | 1.18 | 1.26 | 1.02 | 0.98 | 0.17 | 0.196 | 0.8582 | 0.36 |
|  | 5 | 1.03 | 1.06 | 0.998 | 1.07 | 1.03 | 0.97 | 0.1725 | 0.194 | 0.8267 | 0.33 |
|  | 10 | 0.77 | 0.83 | 0.73 | 0.87 | 1.05 | 0.95 | 0.18 | 0.19 | 0.7679 | 0.27 |
| 8 | 3 | 1.32 | 1.34 | 1.30 | 1.36 | 1.02 | 0.98 | 0.13 | 0.20 | 0.8782 | 0.42 |
|  | 4 | 0.99 | 1.03 | 0.97 | 1.07 | 1.03 | 0.97 | 0.13 | 0.19 | 0.8247 | 0.36 |
|  | 5 | 0.86 | 0.90 | 0.82 | 0.94 | 1.05 | 0.95 | 0.13 | 0.19 | 0.7915 | 0.33 |
|  | 10 | 0.64 | 0.71 | 0.60 | 0.76 | 1.07 | 0.93 | 0.13 | 0.19 | 0.7333 | 0.27 |
| 10 | 3 | 1.18 | 1.20 | 1.15 | 1.23 | 1.03 | 0.97 | 0.10 | 0.19 | 0.8613 | 0.42 |
|  | 4 | 0.90 | 0.94 | 0.86 | 0.98 | 1.04 | 0.96 | 0.10 | 0.19 | 0.8060 | 0.36 |
|  | 5 | 0.77 | 0.82 | 0.73 | 0.87 | 1.05 | 0.95 | 0.10 | 0.19 | 0.7726 | 0.33 |
|  | 10 | 0.57 | 0.65 | 0.53 | 0.71 | 1.08 | 0.92 | 0.11 | 0.18 | 0.7168 | 0.27 |

Table 13 Numerical results for the model with switch-over times, $\lambda=3, \mu_{2}=5, p_{1}=p_{2}=0.5, \gamma_{2}=3$

| $\mu_{1}$ | $\gamma_{1}$ | $\mathbb{E}\left[L_{1}\right]$ | $\mathbb{E}\left[L_{2}\right]$ | $\mathbb{E}\left[W_{1}\right]$ | $\mathbb{E}\left[W_{2}\right]$ | $\lambda_{e f f}^{1}$ | $\lambda_{e f f}^{2}$ | $\rho_{e f f}^{1}$ | $\rho_{e f f}^{2}$ | $\operatorname{Cor}\left(L_{1}, L_{2}\right)$ | $\mathbb{P}($ Switch $)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 5 | 10 | 37.62 | 37.67 | 23.83 | 26.50 | 1.58 | 1.42 | 0.32 | 0.28 | 0.9997 | 0.39 |
|  | 12 | 15.81 | 15.87 | 9.97 | 11.23 | 1.41 | 1.59 | 0.32 | 0.28 | 0.9985 | 0.37 |
|  | 14 | 11.11 | 11.18 | 6.98 | 7.94 | 1.41 | 1.59 | 0.32 | 0.28 | 0.9971 | 0.36 |
| 6 | 10 | 6.64 | 6.71 | 4.18 | 4.76 | 1.59 | 1.41 | 0.27 | 0.28 | 0.9921 | 0.39 |
|  | 12 | 5.22 | 5.30 | 3.27 | 3.78 | 1.60 | 1.40 | 0.27 | 0.28 | 0.9878 | 0.38 |
|  | 14 | 4.52 | 4.59 | 2.81 | 3.29 | 1.61 | 1.39 | 0.27 | 0.28 | 0.9842 | 0.37 |
| 8 | 5 | 14.11 | 14.16 | 9.00 | 9.88 | 1.57 | 1.43 | 0.19 | 0.27 | 0.9981 | 0.49 |
|  | 8 | 3.91 | 3.97 | 2.44 | 2.84 | 1.60 | 1.40 | 0.20 | 0.28 | 0.9793 | 0.42 |
|  | 10 | 3.08 | 3.16 | 1.91 | 2.27 | 1.61 | 1.39 | 0.20 | 0.28 | 0.9692 | 0.39 |
|  | 12 | 2.69 | 2.77 | 1.66 | 2.00 | 1.62 | 1.38 | 0.20 | 0.28 | 0.9616 | 0.38 |
|  | 14 | 2.46 | 2.54 | 1.51 | 1.85 | 1.63 | 1.37 | 0.20 | 0.27 | 0.9559 | 0.37 |
| 10 | 4 | 21.44 | 21.48 | 13.75 | 14.91 | 1.56 | 1.44 | 0.16 | 0.29 | 0.9992 | 0.53 |
|  | 5 | 6.15 | 6.20 | 3.89 | 4.36 | 1.58 | 1.42 | 0.16 | 0.28 | 0.9908 | 0.49 |
|  | 8 | 2.74 | 2.81 | 1.70 | 2.03 | 1.61 | 1.39 | 0.16 | 0.28 | 0.9627 | 0.42 |
|  | 10 | 2.27 | 2.35 | 1.39 | 1.71 | 1.62 | 1.38 | 0.16 | 0.27 | 0.9501 | 0.39 |
|  | 12 | 2.03 | 2.11 | 1.24 | 1.55 | 1.64 | 1.36 | 0.16 | 0.27 | 0.9412 | 0.38 |

2. In both tables, as $\mu_{1}$ increases, the effective arrival rate to $Q_{1}$, i.e. $\lambda_{e f f}^{1}$, increases. Clearly, a fast service rate in $Q_{1}$ shortens the queue length, and therefore rises the probability that a newly arrival will join $Q_{1}$.
3. The correlation coefficient between $L_{1}$ and $L_{2}$ is always positive, as obtained in Tables 1, 2, 3, 4 and 5 . This again follows from the JSQ policy, as explained in Sect. 6.1.
4. The proportion of time the server is switching between queues is a decreasing function of $\gamma_{1}$. However, it is less sensitive to the service rate $\mu_{1}$.

## 7 Concluding remarks

The combined operating policy 'Join the Shortest Queue-Serve the Longest Queue' is analyzed under the non-preemptive service regime for a 2-queue Markovian system attended by a single server. The system is formulated in an innovative way, where, instead of defining an un-bounded 2-dimensional state space ( $L_{1}, L_{2}$ ), where $L_{i}$ represents the number of customers in $Q_{i}, i=1,2$, the system is characterized by the couple $L_{1}$ and $D=L_{1}-L_{2}$. This leads to a 2-dimensional state space with finite, and small, dimension for $D$, and infinite dimension only for $L_{1}$. The resulting QBD process enables the combined use of PGFs method and matrix geometric analysis. The results of the non-preemptive service regime are compared numerically with the corresponding results of its twin preemptive service regime in numerous tables for a wide range of parameter values. Among many insights, it is shown that when the overall traffic intensity $\rho=\rho_{e f f}^{1}+\rho_{e f f}^{2}$ approaches 1 , and $p_{1} \leq p_{2}$, the values of $\mathbb{E}\left[L_{i}\right]$ and $\mathbb{E}\left[W_{i}\right]$ under the preemptive regime are smaller than their corresponding values under the non-preemptive discipline. The ratio changes when $p_{1}>p_{2}$. When $\rho$ is small, the differences between the performance measures under the two regimes are small. In terms of total operating cost (customers' sojourn times and server's switches), it is shown that there are cases where the non-preemptive regime is more efficient economically, while in other cases the preemptive regime is preferable.

## Declarations

Data availability statements We do not analyze or generate any datasets, because our work proceeds within a theoretical and mathematical approach.

Conflict of interest All authors certify that they have no affiliations with or involvement in any organization or entity with any financial interest or non-financial interest in the subject matter or materials discussed in this manuscript.

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